

Part 1: From Physics to Statistics

or

“What quantity is reconstructed?”
(in emission tomography)

Outline

- Decay phenomena and fundamental assumptions
- Idealized detectors
- Random phenomena
- Poisson measurement statistics
- State emission tomography reconstruction problem

What Object is Reconstructed?

In *emission imaging*, our aim is to image the *radiotracer distribution*.

The what?

At time $t = 0$ we inject the patient with some *radiotracer*, containing a “large” number N of metastable atoms of some radionuclide.

Let $\vec{X}_k(t) \in \mathbb{R}^3$ denote the position of the k th *tracer atom* at time t .

These positions are influenced by blood flow, patient physiology, and other unpredictable phenomena such as Brownian motion.

The ultimate imaging device would provide an exact list of the spatial locations $\vec{X}_1(t), \dots, \vec{X}_N(t)$ of all tracer atoms for the entire scan.

Would this be enough?

Atom Positions or Statistical Distribution?

Repeating a scan would yield different tracer atom sample paths $\left\{ \vec{X}_k(t) \right\}_{k=1}^N$.

∴ statistical formulation

Assumption 1. The spatial locations of individual tracer atoms at any time $t \geq 0$ are *independent* random variables that are all *identically distributed* according to a common probability density function (pdf) $f_{\vec{X}(t)}(\vec{x})$.

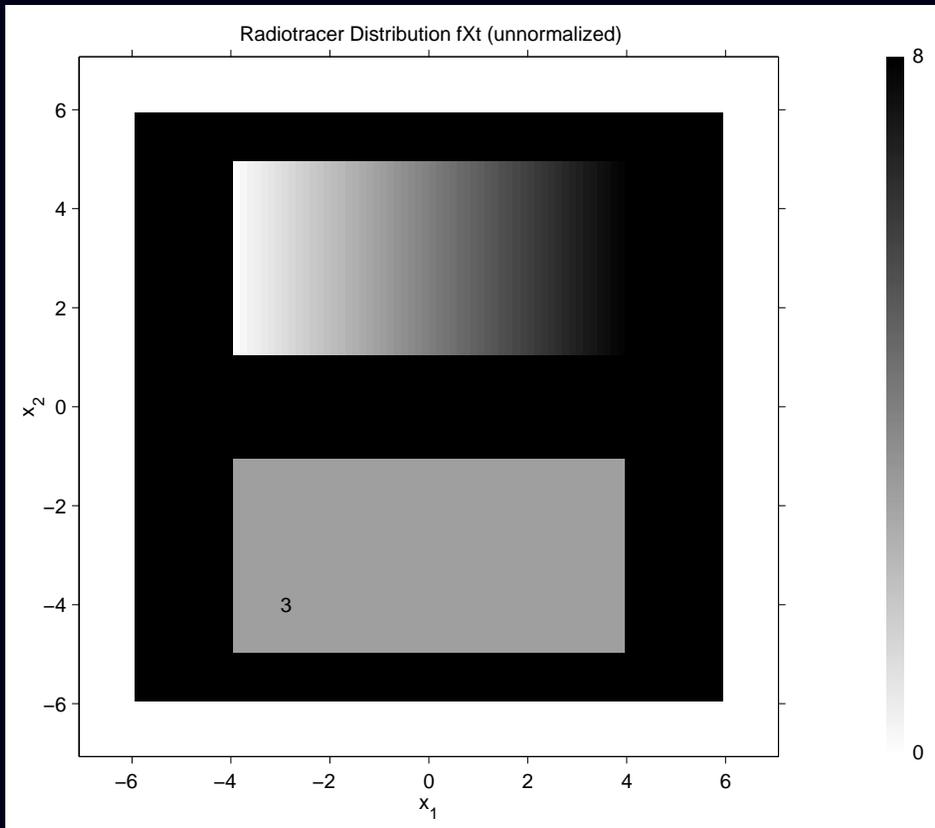
This pdf is determined by patient physiology and tracer properties.

Larger values of $f_{\vec{X}(t)}(\vec{x})$ correspond to “hot spots” where the tracer atoms tend to be located at time t . Units: inverse volume, e.g., atoms per cubic centimeter.

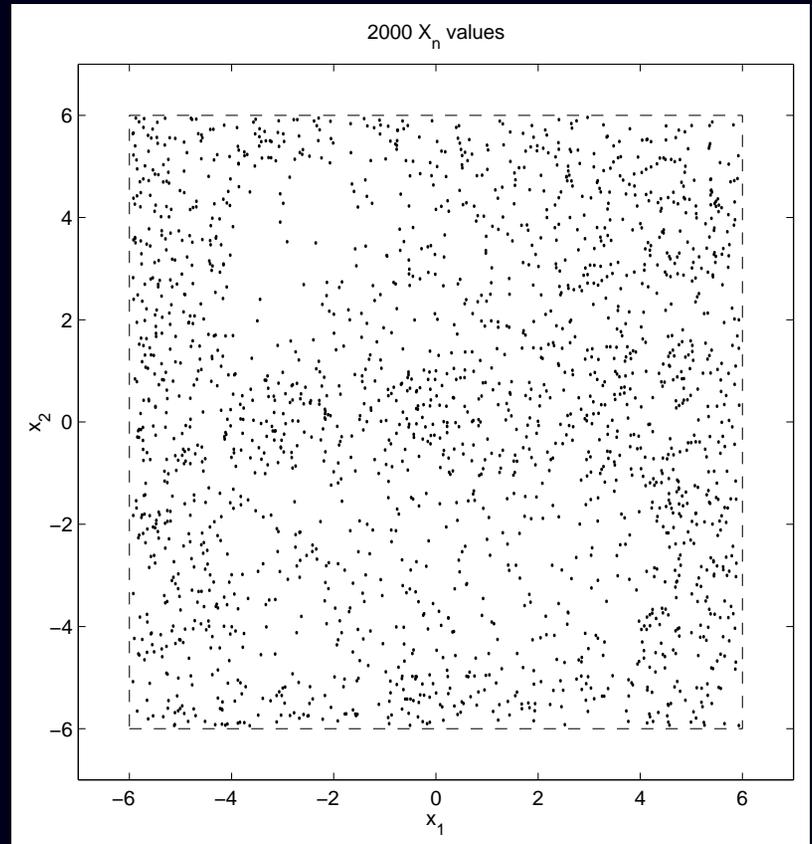
The *radiotracer distribution* $f_{\vec{X}(t)}(\vec{x})$ is the quantity of interest.

(Not $\left\{ \vec{X}_k(t) \right\}_{k=1}^N$!)

Example: Perfect Detector



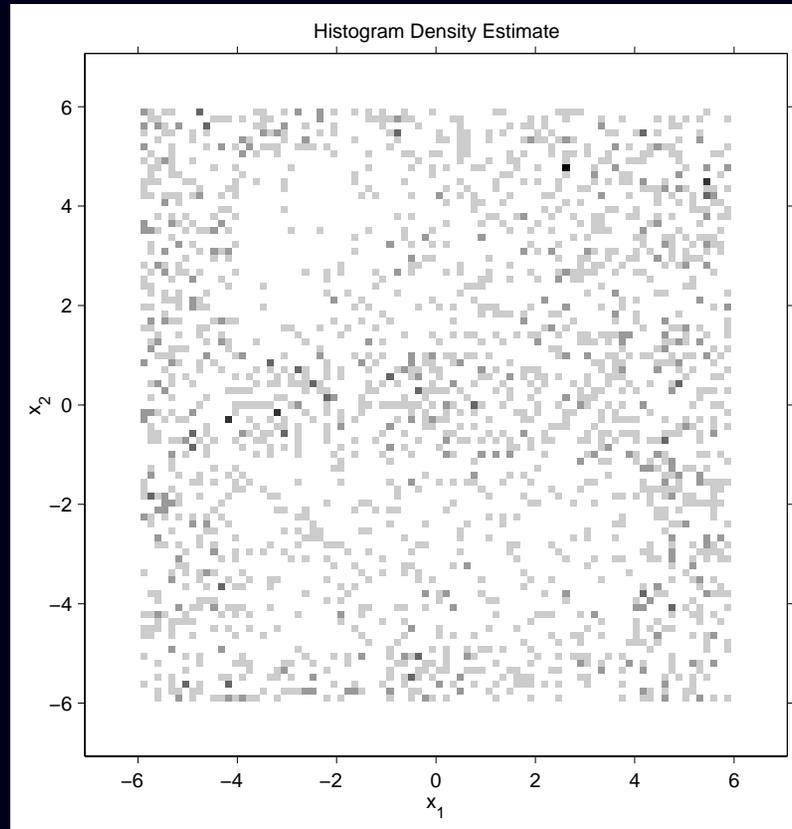
True radiotracer distribution $f_{\vec{X}(t)}(\vec{x})$
at some time t .



A realization of $N = 2000$ i.i.d.
atom positions (dots) recorded
"exactly."

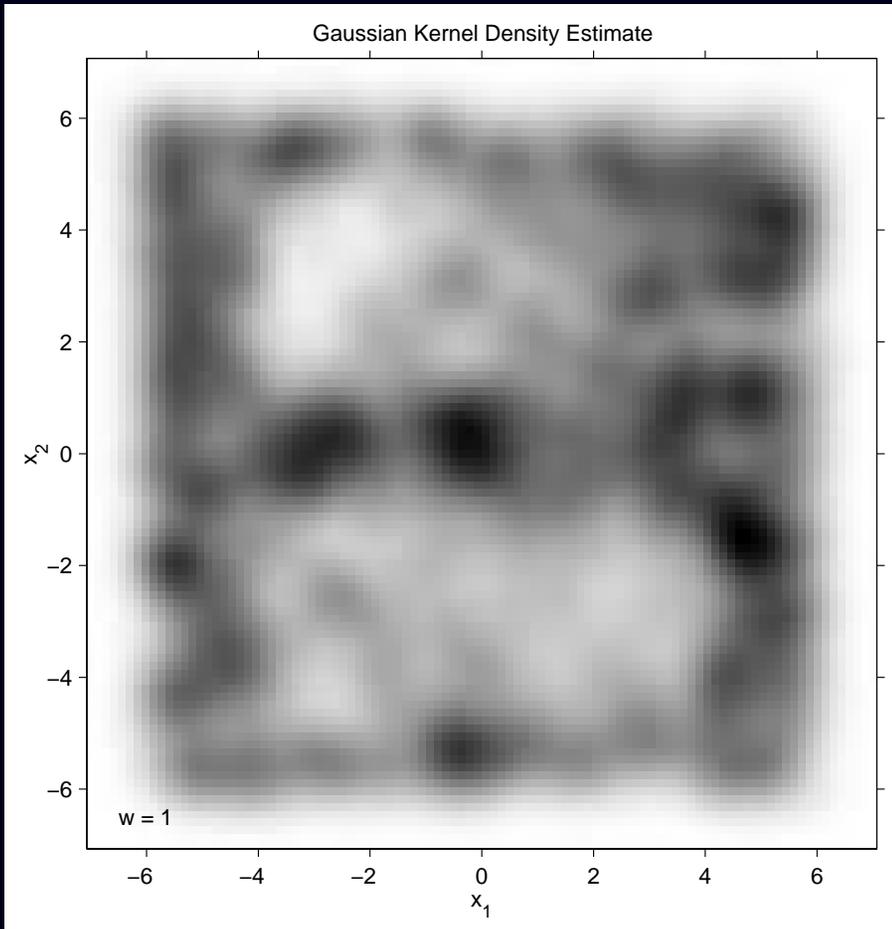
Little similarity!

Binning/Histogram Density Estimator

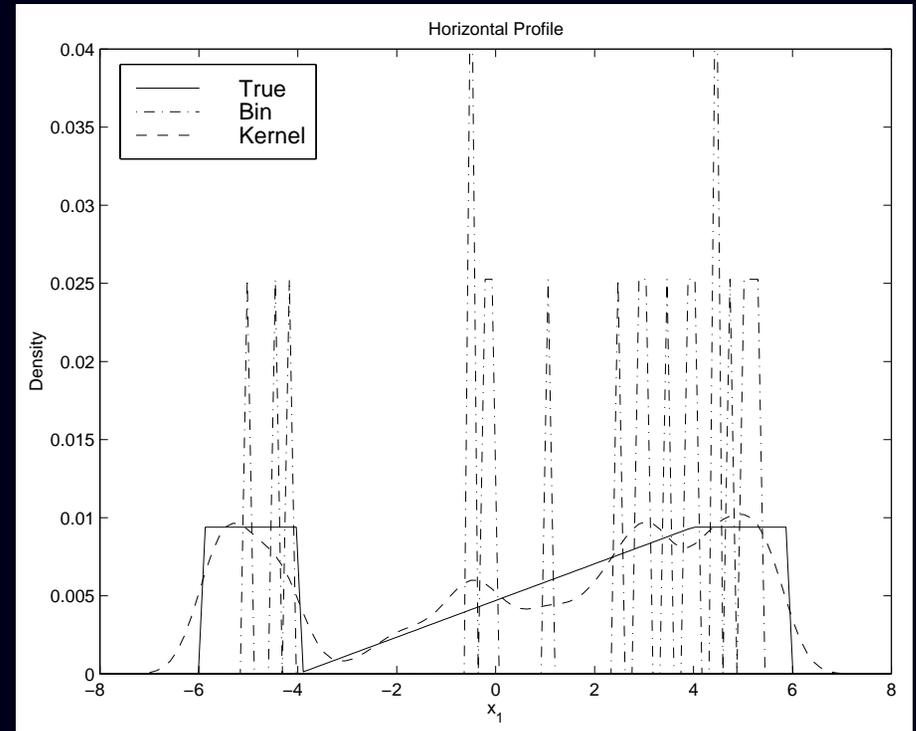


Estimate of $f_{\vec{X}(t)}(\vec{x})$ formed by histogram binning of $N = 2000$ points.
Ramp remains difficult to visualize.

Kernel Density Estimator



Gaussian kernel density estimator for $f_{\vec{X}(t)}(\vec{x})$ from $N = 2000$ points.



Horizontal profiles at $x_2 = 3$ through density estimates.

Poisson Spatial Point Process

Assumption 2. The number of injected tracer atoms N has a Poisson distribution with some mean

$$\mu_N \triangleq E[N] = \sum_{n=0}^{\infty} nP[N = n].$$

Let $N(B)$ denote the number of tracer atoms that have spatial locations in any set $B \subset \mathbb{R}^3$ (VOI) at time t_0 after injection.

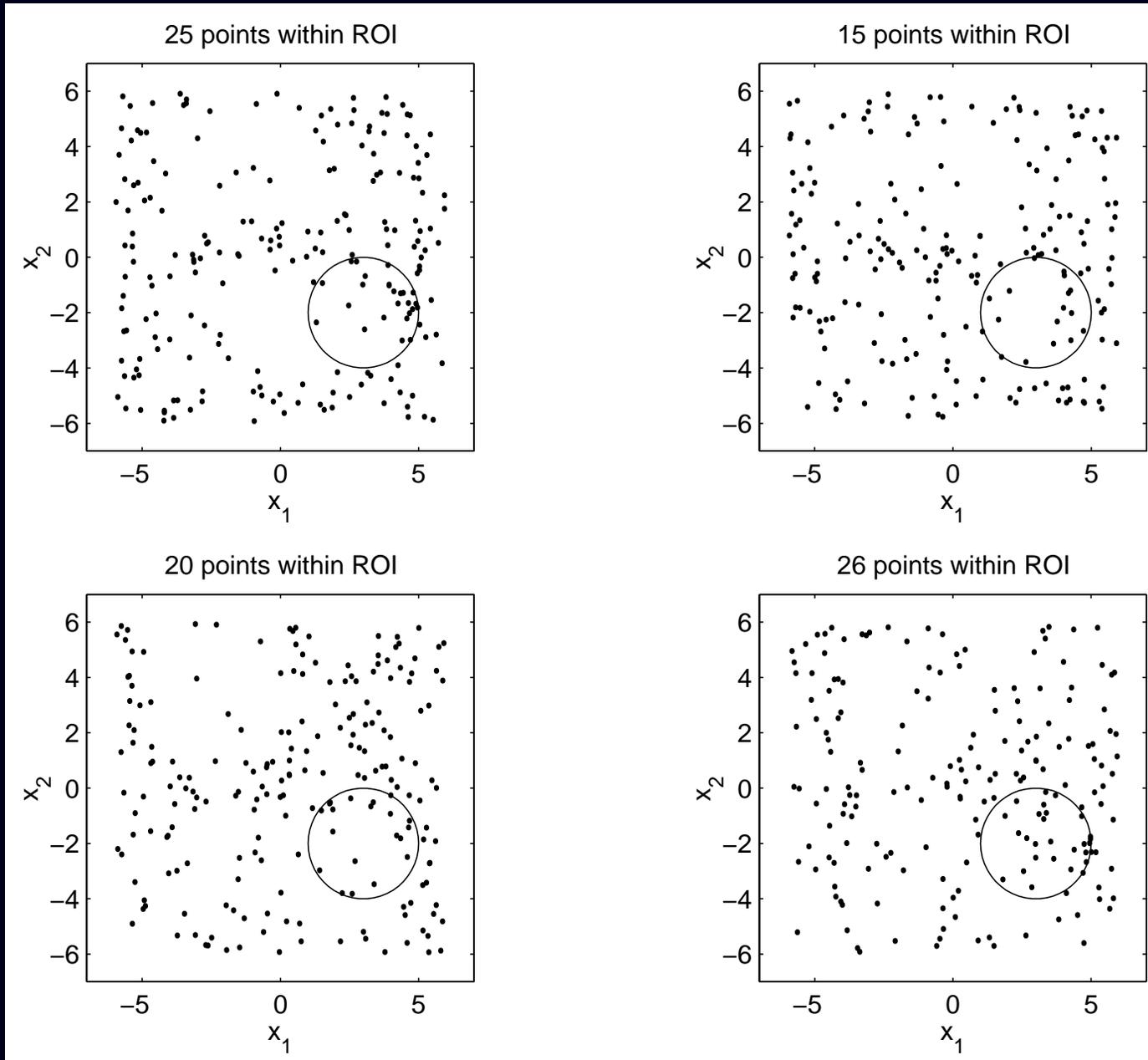
$N(\cdot)$ is called a *Poisson spatial point process*.

Fact. For any set B , $N(B)$ is Poisson distributed with mean:

$$E[N(B)] = E[N]P[\vec{X} \in B] = \mu_N \int_B f_{\vec{X}(t_0)}(\vec{x}) d\vec{x}.$$

Poisson N injected atoms + i.i.d. locations \Rightarrow Poisson point process

Illustration of Point Process ($\mu_N = 200$)



Radionuclide Decay

Preceding quantities are all unobservable.

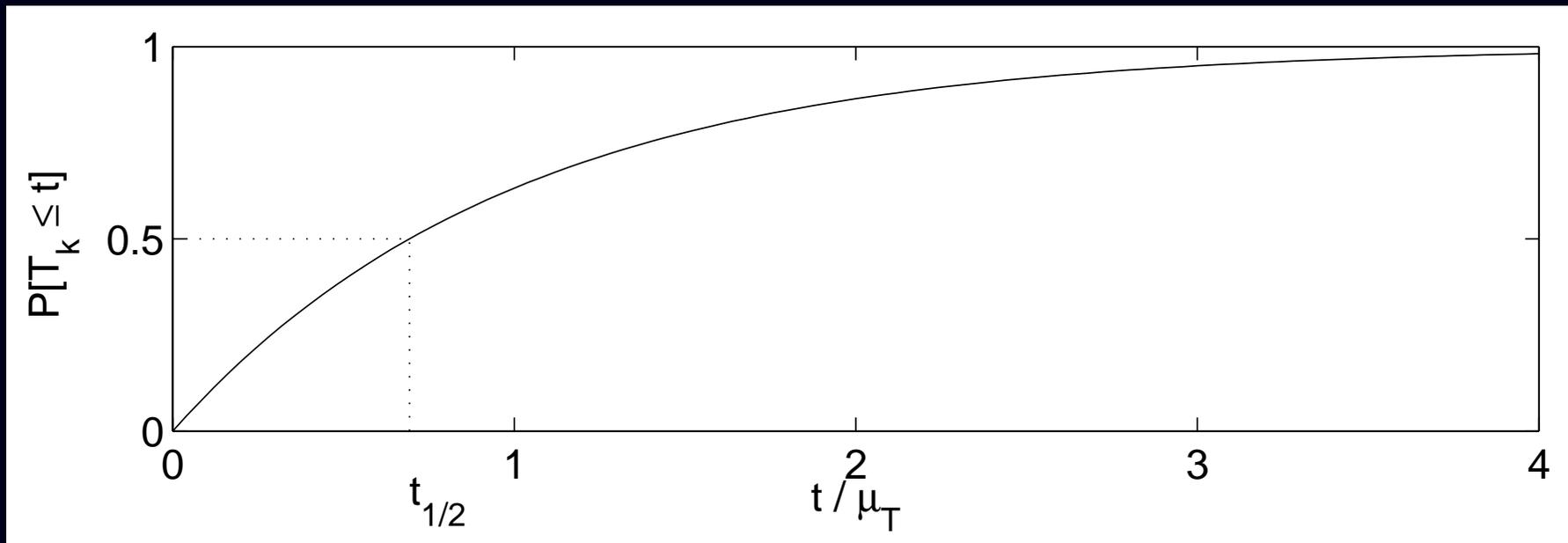
We “observe” a tracer atom only when it decays and emits photon(s).

The time that the k th tracer atom decays is a random variable T_k .

Assumption 3. The T_k 's are statistically *independent* random variables, and are independent of the (random) spatial location.

Assumption 4. Each T_k has an exponential distribution with mean $\mu_T = t_{1/2}/\ln 2$.

Cumulative distribution function: $P[T_k \leq t] = 1 - \exp(-t/\mu_T)$



Statistics of an Ideal Decay Counter

Let $K(t, B)$ denote the number of tracer atoms that decay by time t , and that were located in the VOI $B \subset \mathbb{R}^3$ at the time of decay.

Fact. $K(t, B)$ is a *Poisson counting process* with mean

$$E[K(t, B)] = \int_0^t \int_B \lambda(\vec{x}, \tau) d\vec{x} d\tau,$$

where the (nonuniform) *emission rate density* is given by

$$\lambda(\vec{x}, t) \triangleq \mu_N \frac{e^{-t/\mu_T}}{\mu_T} \cdot f_{\vec{X}(t)}(\vec{x}).$$

Ingredients: “dose,” “decay,” “distribution”

Units: “counts” per unit time per unit volume, e.g., $\mu\text{Ci/cc}$.

“Photon emission is a Poisson process”

What about the actual measurement statistics?

Idealized Detector Units

A nuclear imaging system consists of n_d conceptual *detector units*.

Assumption 5. Each decay of a tracer atom produces a recorded count in at most one detector unit.

Let $S_k \in \{0, 1, \dots, n_d\}$ denote the index of the incremented detector unit for decay of k th tracer atom. ($S_k = 0$ if decay is undetected.)

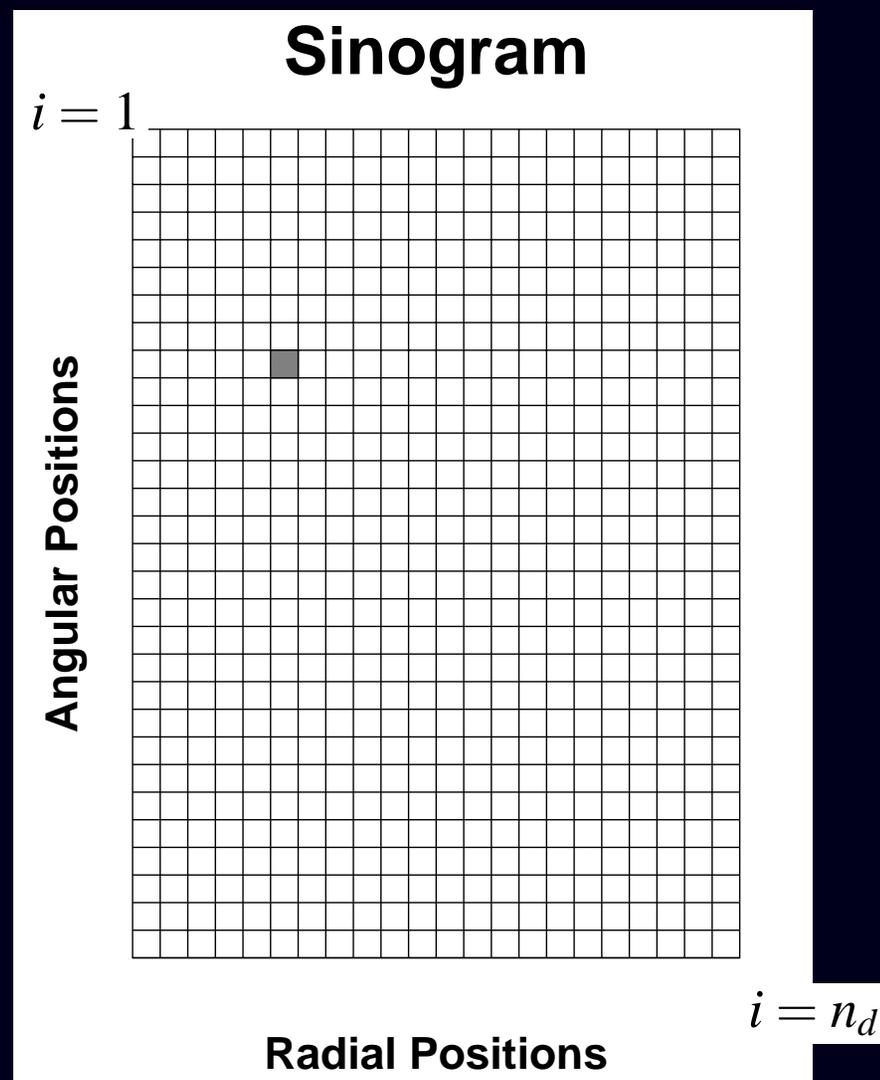
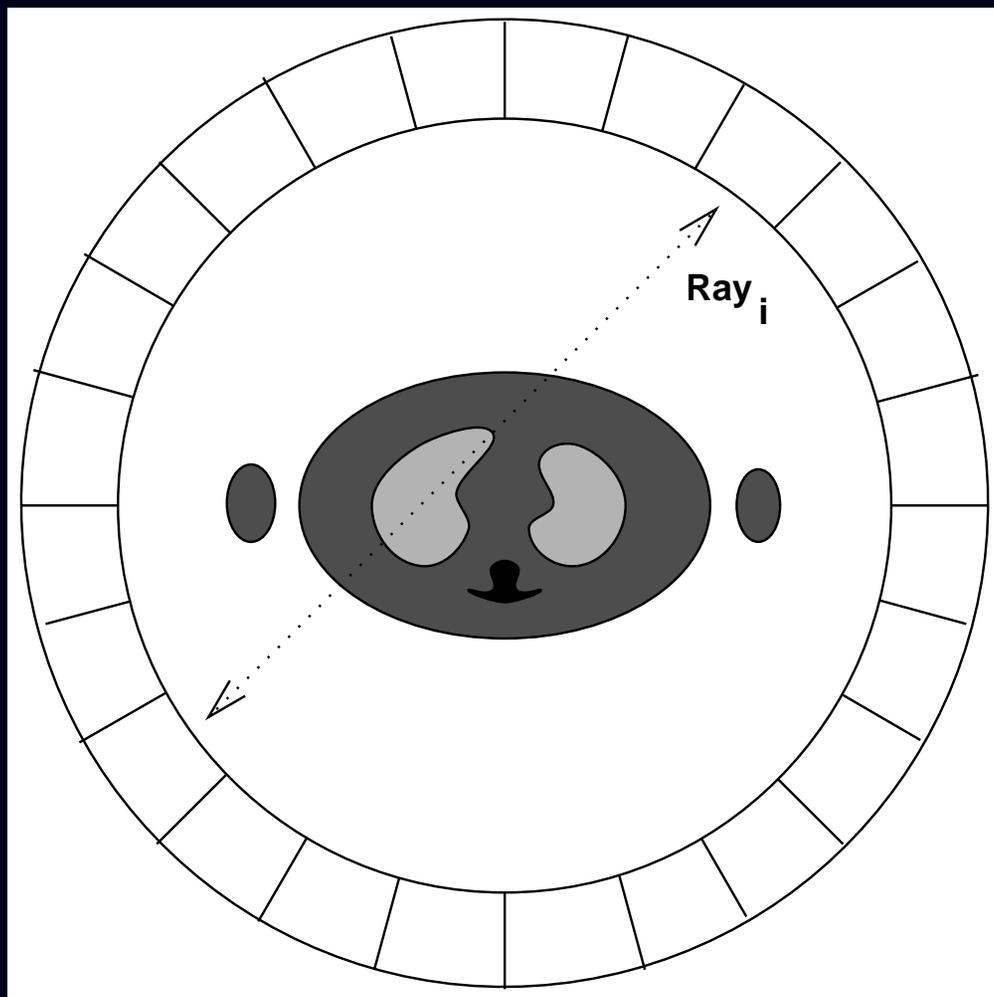
Assumption 6. The S_k 's satisfy the following conditional independence:

$$P\left(S_1, \dots, S_N \mid N, T_1, \dots, T_N, \vec{X}_1(\cdot), \dots, \vec{X}_N(\cdot)\right) = \prod_{k=1}^N P\left(S_k \mid \vec{X}_k(T_k)\right).$$

The recorded bin for the k th tracer atom's decay depends only on its position when it decays, and is independent of all other tracer atoms.

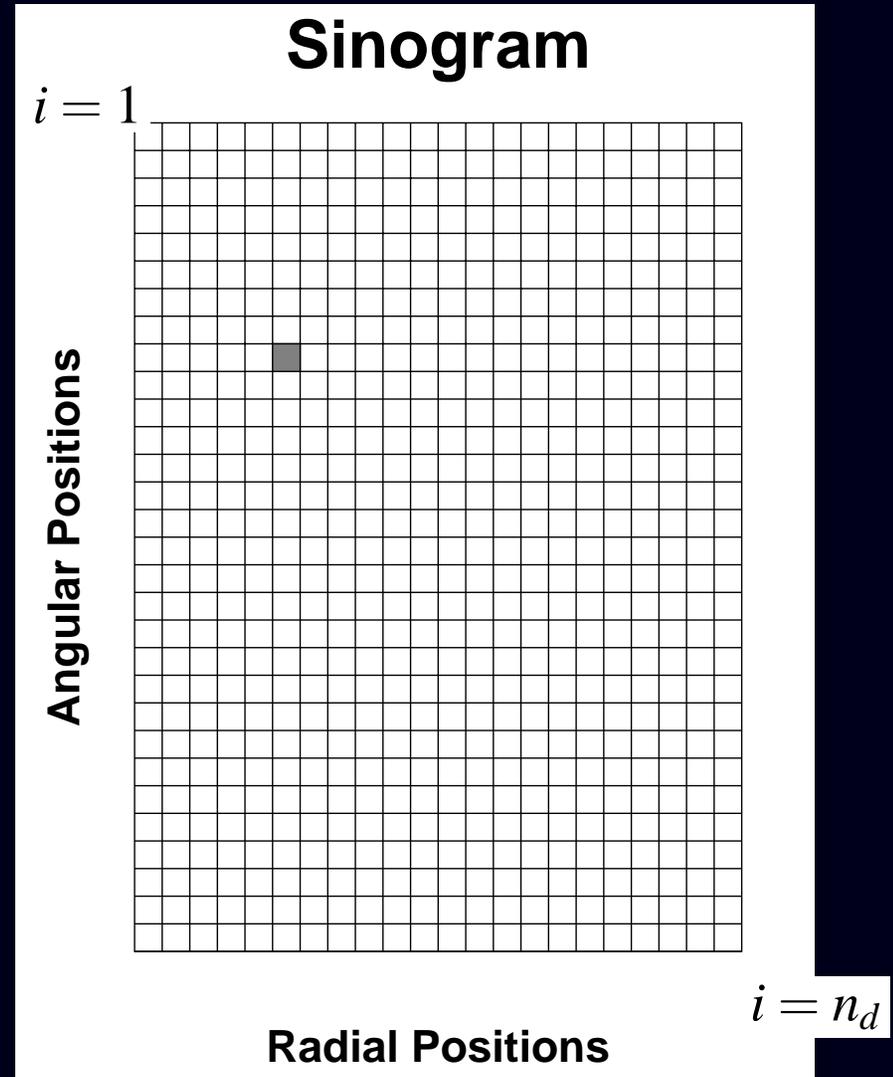
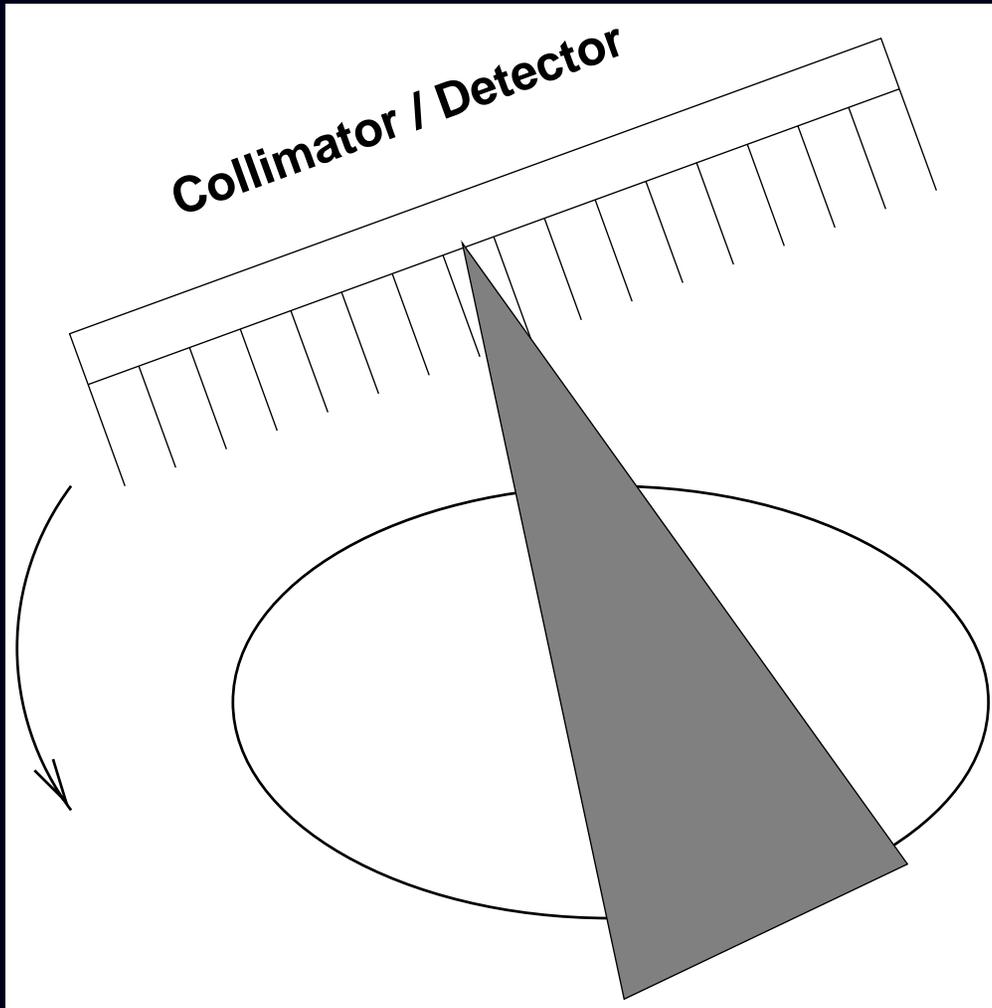
(No event pileup; no deadtime losses.)

PET Example



$$n_d \leq (n_{\text{crystals}} - 1) \cdot n_{\text{crystals}} / 2$$

SPECT Example



$$n_d = n_{\text{radial_bins}} \cdot n_{\text{angular_steps}}$$

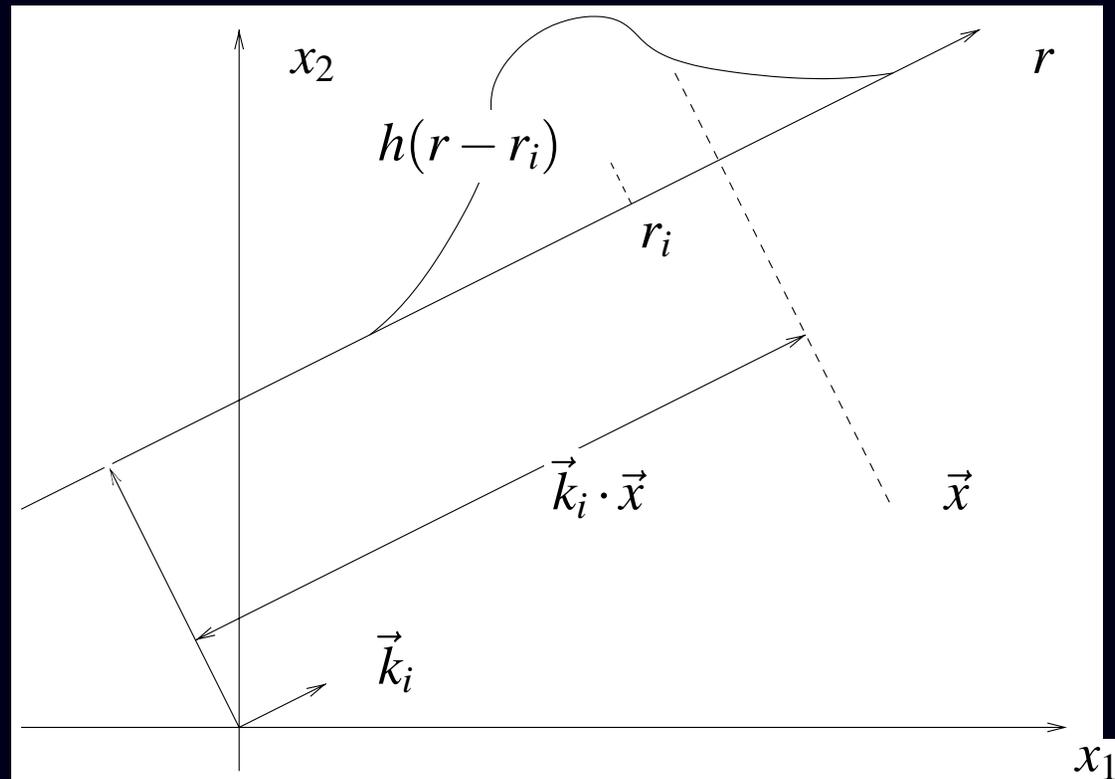
Detector Unit Sensitivity Patterns

Spatial localization:

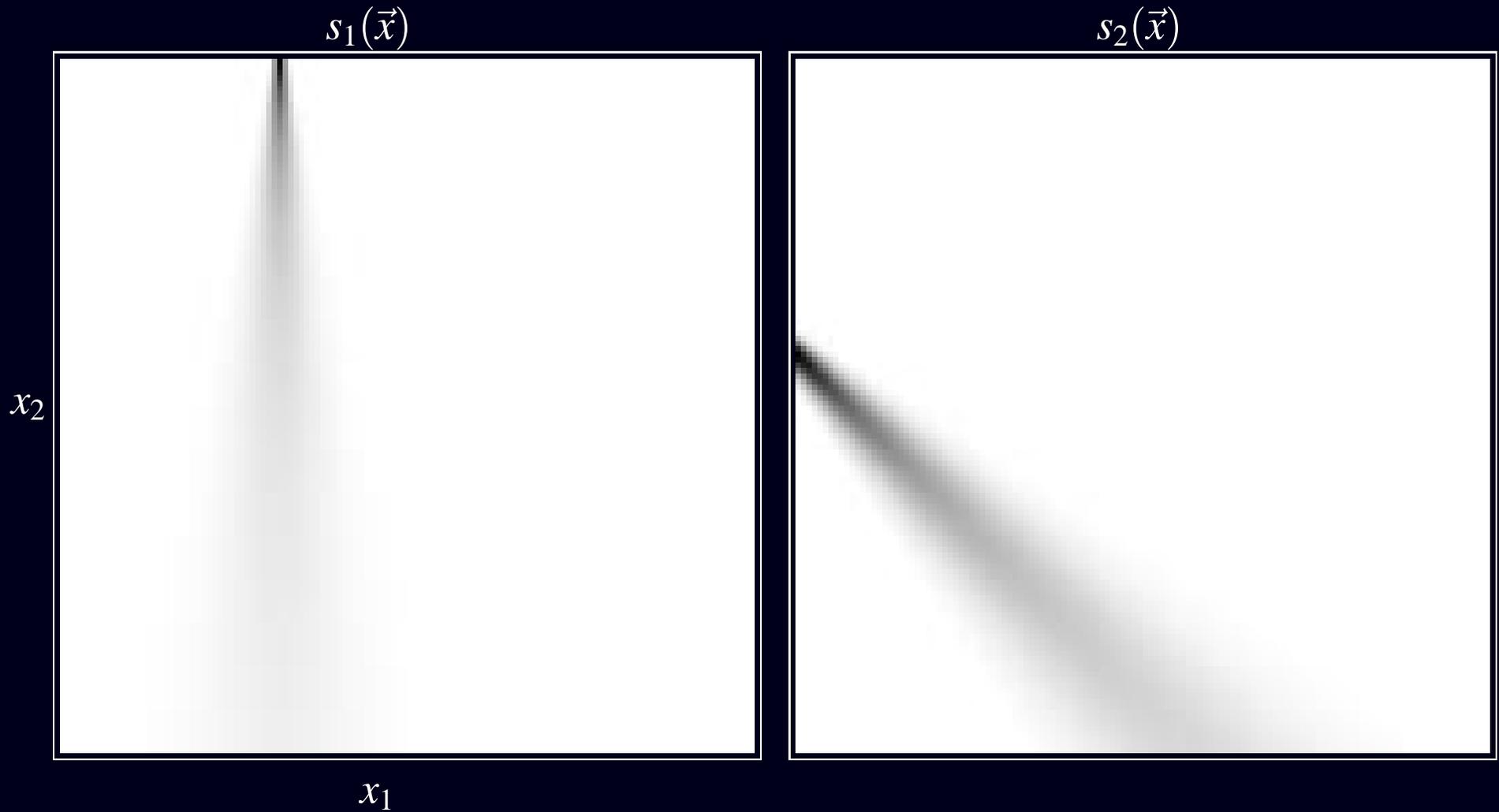
$s_i(\vec{x}) \triangleq$ probability that decay at \vec{x} is recorded by i th detector unit.

Idealized Example. Shift-invariant PSF: $s_i(\vec{x}) = h(\vec{k}_i \cdot \vec{x} - r_i)$

- r_i is the radial position of i th ray
- \vec{k}_i is the unit vector orthogonal to i th parallel ray
- $h(\cdot)$ is the shift-invariant radial PSF (e.g., Gaussian bell or rectangular function)

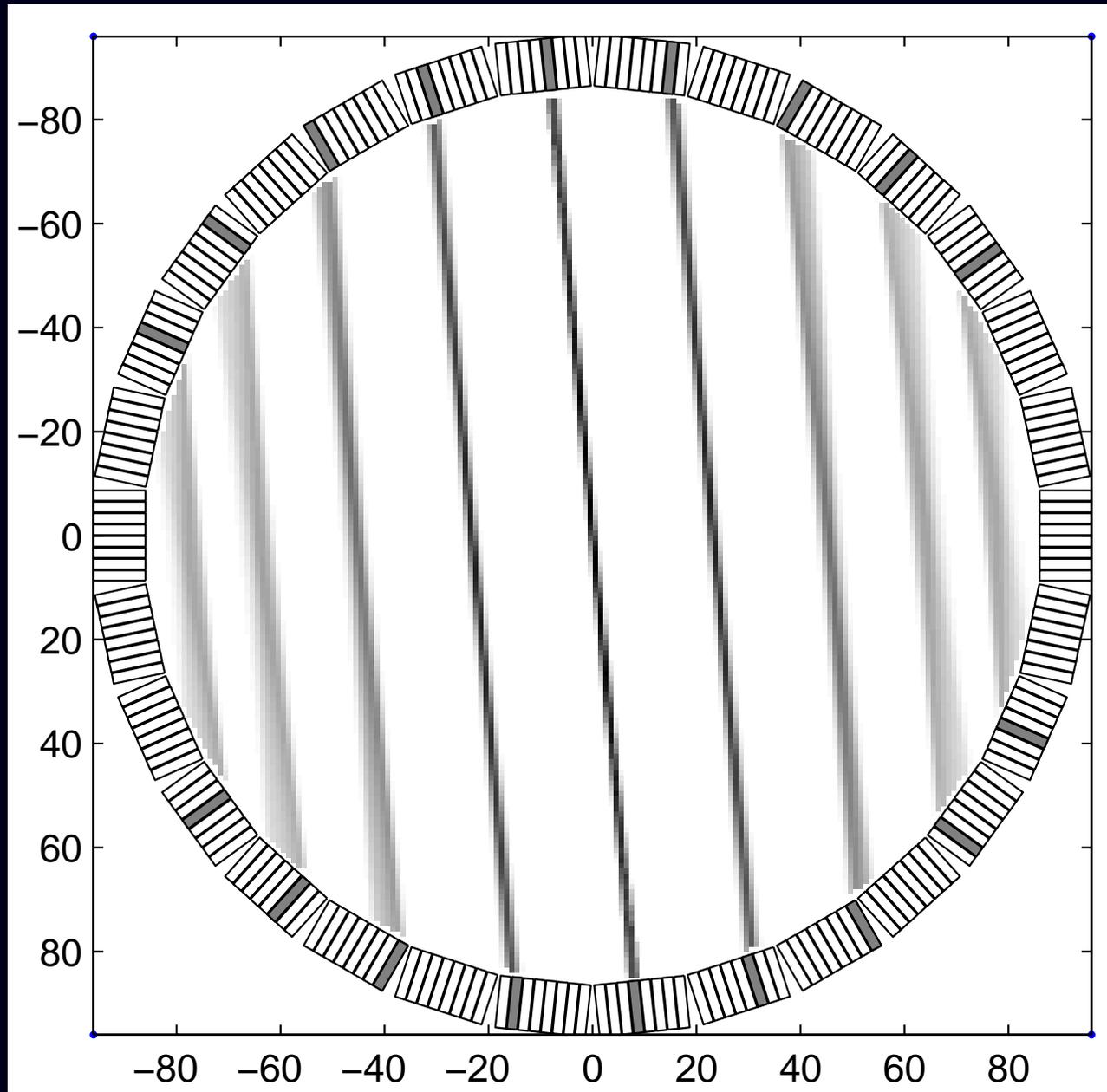


Example: SPECT Detector-Unit Sensitivity Patterns



Two representative $s_i(\vec{x})$ functions for a collimated Anger camera.

Example: PET Detector-Unit Sensitivity Patterns



Detector Unit Sensitivity Patterns

$s_i(\vec{x})$ can include the effects of

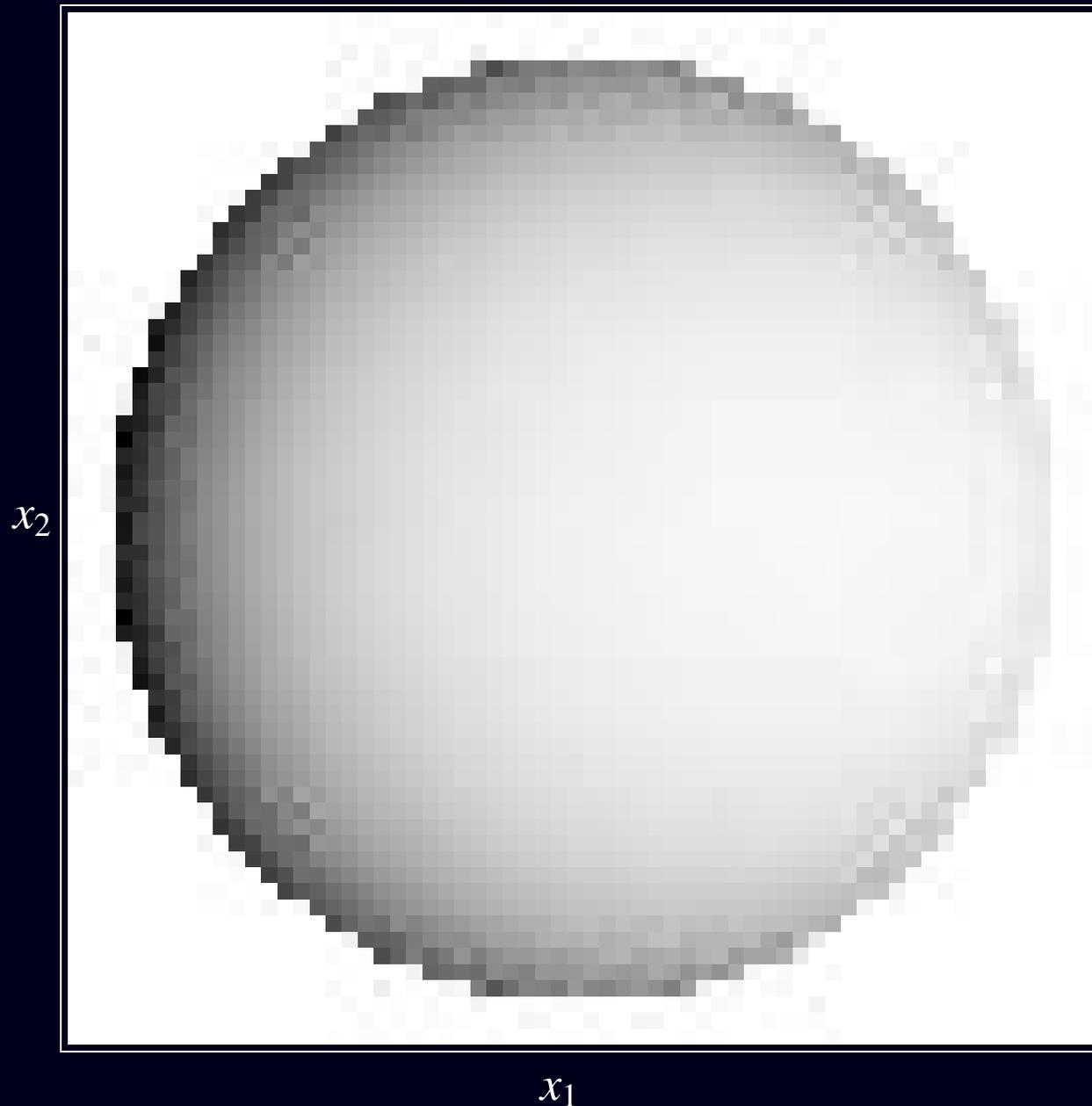
- geometry / solid angle
- collimation
- scatter
- attenuation
- detector response / scan geometry
- duty cycle (dwell time at each angle)
- detector efficiency
- positron range, noncollinearity
- ...

System sensitivity pattern:

$$s(\vec{x}) \triangleq \sum_{i=1}^{n_d} s_i(\vec{x}) = 1 - s_0(\vec{x}) \leq 1$$

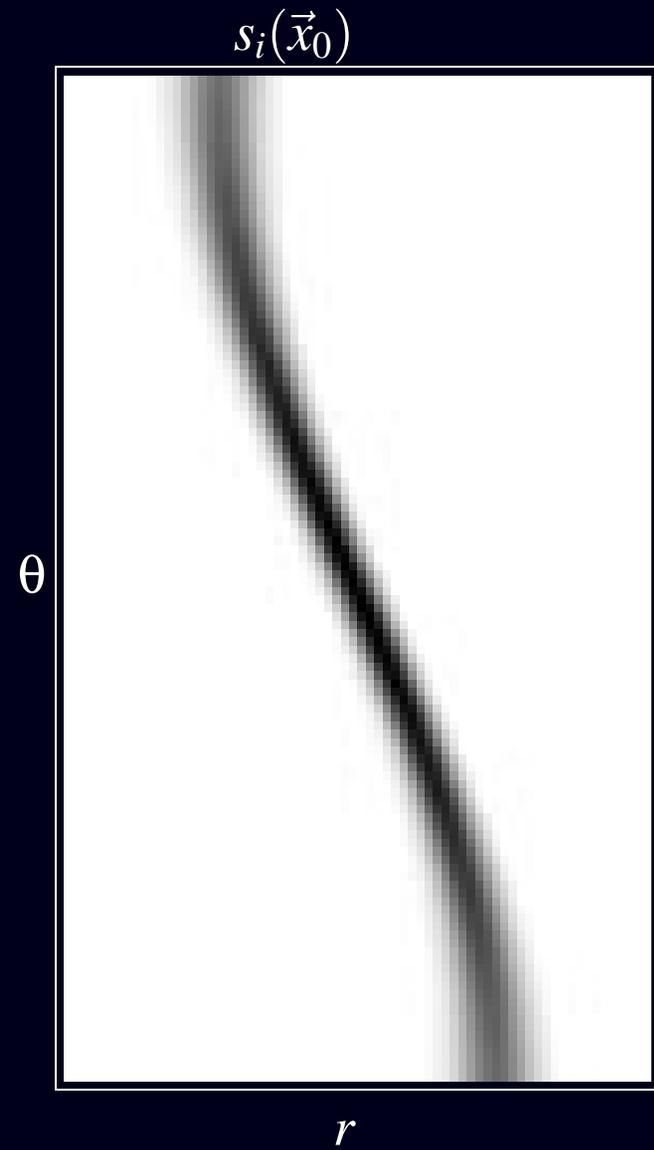
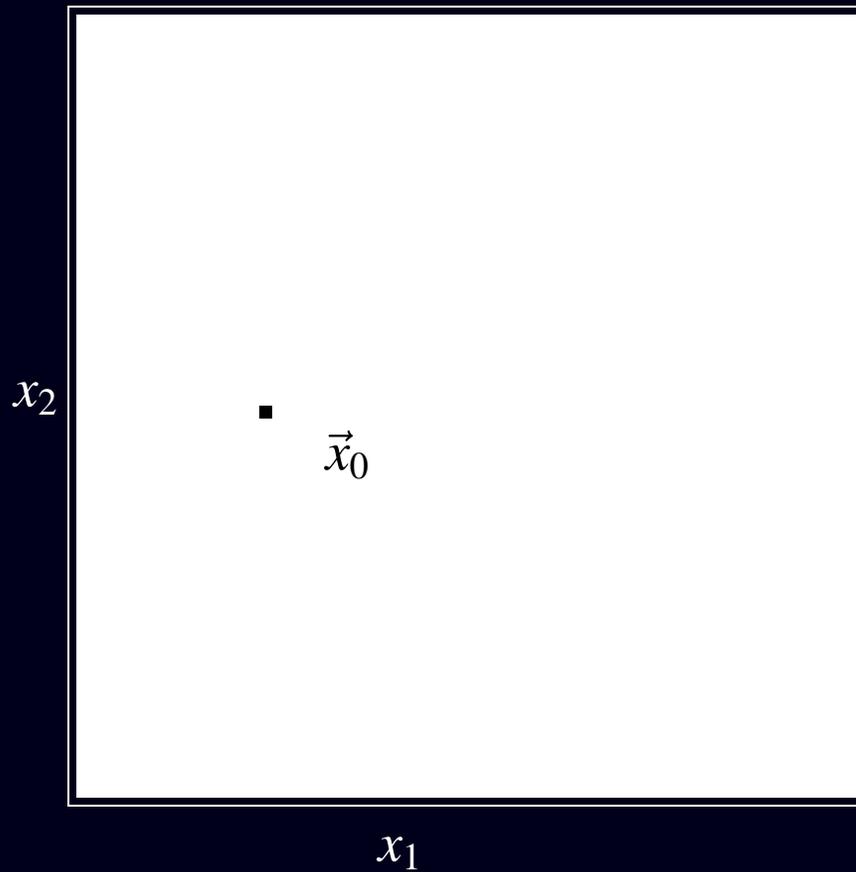
(probability that decay at location \vec{x} will be detected *at all* by system)

System Sensitivity Pattern $s(\vec{x})$



Example: collimated 180° SPECT system with uniform attenuation.

Detection Probabilities $s_i(\vec{x}_0)$ (vs det. unit index i)



Summary of Random Phenomena

- Number of tracer atoms injected N
- Spatial locations of tracer atoms $\{\vec{X}_k\}_{k=1}^N$
- Time of decay of tracer atoms $\{T_k\}_{k=1}^N$
- Detection of photon $[S_k \neq 0]$
- Recording detector unit $\{S_k\}_{i=1}^{n_d}$

Emission Scan

Record events in each detector unit for $t_1 \leq t \leq t_2$.

$Y_i \triangleq$ number of events recorded by i th detector unit during scan, for $i = 1, \dots, n_d$.

$$Y_i \triangleq \sum_{k=1}^N 1_{\{S_k=i, T_k \in [t_1, t_2]\}}.$$

The collection $\{Y_i : i = 1, \dots, n_d\}$ is our *sinogram*.

Note $0 \leq Y_i \leq N$.

Fact. Under Assumptions 1-6 above,

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} \right\} \quad (\text{cf "line integral"})$$

and Y_i 's are statistically independent random variables, where the *emission density* is given by

$$\lambda(\vec{x}) = \mu_N \int_{t_1}^{t_2} \frac{1}{\mu_T} e^{-t/\mu_T} f_{\vec{X}(t)}(\vec{x}) dt.$$

(Local number of decays per unit volume during scan.)

Ingredients:

- dose (injected)
- duration of scan
- decay of radionuclide
- distribution of radiotracer

Poisson Statistical Model (Emission)

Actual measured counts = “foreground” counts + “background” counts.

Sources of background counts:

- cosmic radiation / room background
- random coincidences (PET)
- scatter not account for in $s_i(\vec{x})$
- “crosstalk” from transmission sources in simultaneous T/E scans
- anything else not accounted for by $\int s_i(\vec{x})\lambda(\vec{x}) d\vec{x}$

Assumption 7.

The background counts also have independent Poisson distributions.

Statistical model (continuous to discrete)

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x})\lambda(\vec{x}) d\vec{x} + r_i \right\}, \quad i = 1, \dots, n_d$$

r_i : mean number of “background” counts recorded by i th detector unit.

Emission Reconstruction Problem

Estimate the emission density $\lambda(\cdot)$ using (something like) this model:

$$Y_i \sim \text{Poisson} \left\{ \int s_i(\vec{x}) \lambda(\vec{x}) d\vec{x} + r_i \right\}, \quad i = 1, \dots, n_d.$$

Knowns:

- $\{Y_i = y_i\}_{i=1}^{n_d}$: observed counts from each detector unit
- $s_i(\vec{x})$ sensitivity patterns (determined by system models)
- r_i 's : background contributions (determined separately)

Unknown: $\lambda(\vec{x})$

List-mode acquisitions

Recall that conventional sinogram is temporally binned:

$$Y_i \triangleq \sum_{k=1}^N 1_{\{S_k=i, T_k \in [t_1, t_2]\}}.$$

This binning discards temporal information.

List-mode measurements: record all (detector,time) pairs in a list, *i.e.*,

$$\{(S_k, T_k) : k = 1, \dots, N\}.$$

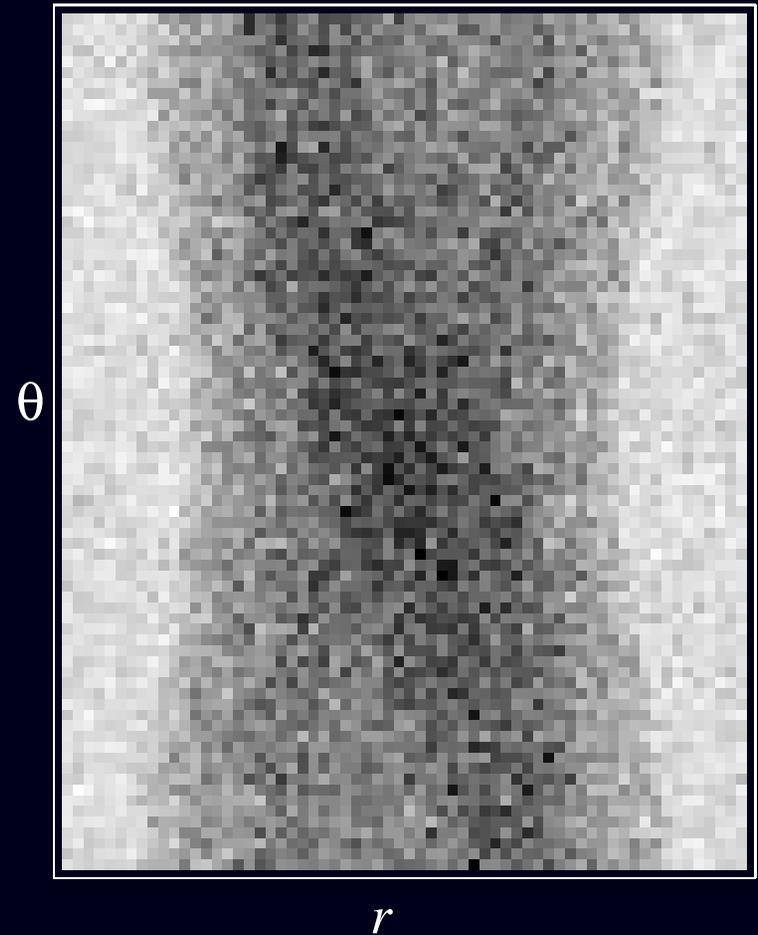
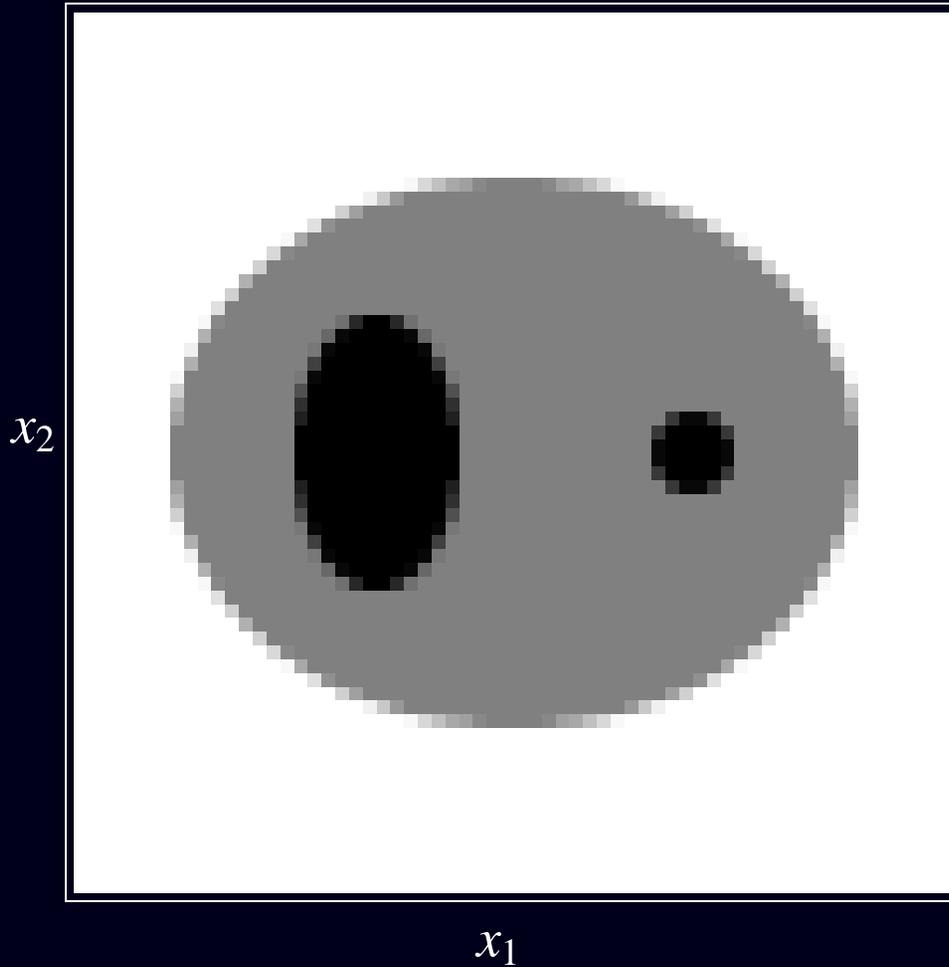
List-mode dynamic reconstruction problem:

$$\text{Estimate } \lambda(\vec{x}, t) \text{ given } \{(S_k, T_k)\}.$$

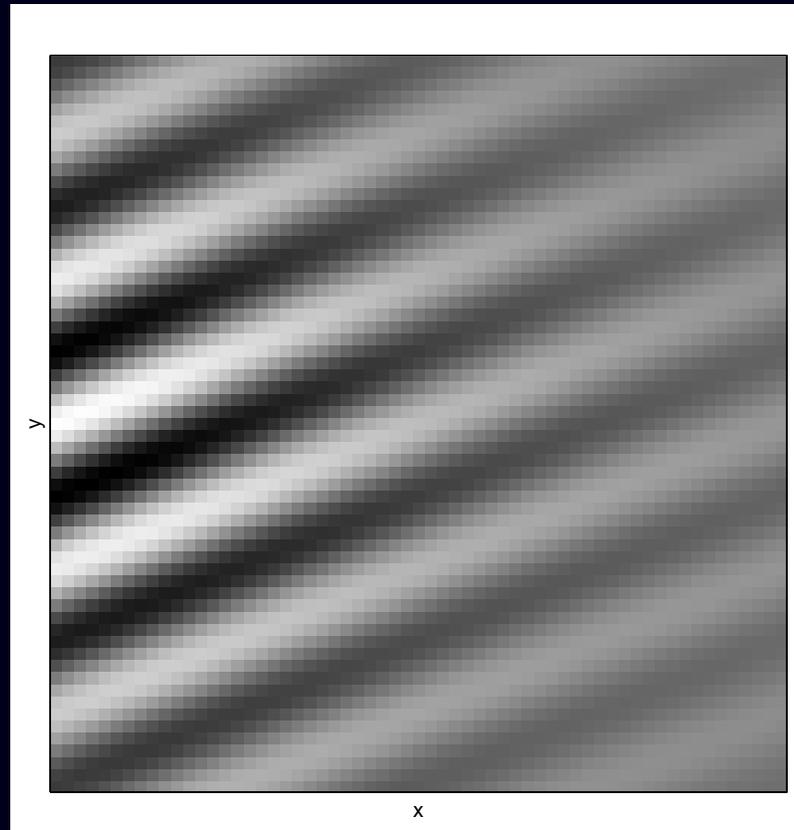
Emission Reconstruction Problem - Illustration

$\lambda(\vec{x})$

$\{Y_i\}$



Example: MRI “Sensitivity Pattern”



Each “k-space sample” corresponds to a sinusoidal pattern weighted by:

- RF receive coil sensitivity pattern
- phase effects of field inhomogeneity
- spin relaxation effects.

$$y_i = \int f(\vec{x}) c_{\text{RF}}(\vec{x}) \exp(-i\omega(\vec{x})t_i) \exp(-t_i/T_2(\vec{x})) \exp\left(-i2\pi\vec{k}(t_i) \cdot \vec{x}\right) d\vec{x} + \epsilon_i$$