

# ADVANCED COMPUTER NETWORKS

Leland et al., "On the Self-Similar Nature of Ethernet Traffic (Extended Version)," *IEEE/ACM Transactions on Networking*, 2(1):1–15, Feb. 1994

# Self-Similarity

Viewing scale not apparent from object appearance

Object features are statistically similar between object parts and the overall object

For example, we always see a jagged line no matter how close we look at a coastline

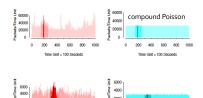
#### References

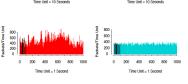
[ENW96] Erramilli, Narayan, and Willinger, "Experimental Queueing Analysis with Long-Range Dependent Packet Traffic," *IEEE/ACM ToN*, 4(2):209–223, Apr. 1996

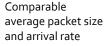
[W+97] Willinger, Taqqu, Sherman, and Wilson, "Self-Similarity Through High-Variability: Statistical Analysis of Ethernet LAN Traffic at the Source Level," *IEEE/ACM ToN*, 5(1):71–86, Feb. 1997

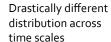
# Ethernet Traffic ...

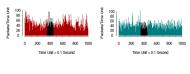
Absence of a natural length of a "burst": at every time scale from msecs to minutes and hours, bursts consist of bursty subperiods separated by *less bursty* subperiods

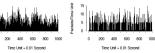












#### A Traffic Trace

Let X be a covariance stationary stochastic process What is:

- a stochastic process:
- a time series of a variable that changes randomly
- a stationary process:
- a time series whose statistical properties: mean, variance, autocorrelation, stay constant over time
- covariance:

by how much two random variables move in tandem

By how much  $X_t$  and  $X_{t+h}$  move in tandem is not a function of time

## Self-Similarity

A traffic trace is self-similar if (equivalently):

1. the variance of the sample mean remains large even as you sample at larger and larger samples (no smoothing out):

$$var(X^{(m)}) \sim a_2 m^{-\beta}, m \to \infty, 0 < \beta < 1, a_2 > 0$$

VS.

$$\operatorname{var}(X^{(m)}) \sim a_{4} m^{-1}, m \to \infty, a_{4} > 0$$

## Sample Mean

Let *X* be a covariance stationary stochastic process

 $X^{(m)}$ , a sample mean, is a new covariance stationary time series obtained by averaging the original series X over non-overlapping blocks of size m

- $X: [1,3,6,2,5,1,8,3,2,\ldots]$
- $X^{(2)}$ :
- $X^{(3)}$ :
- $X^{(4)}$ :

# Self-Similarity

A traffic trace is self-similar if (equivalently):

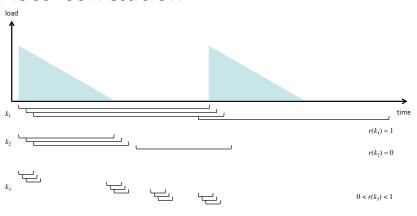
- 2. auto-correlation functions (acf) at various time scales are of the form:  $r(k) \sim k^{-\beta}$ ,  $k \to \infty, 0 < \beta < 1$ 
  - $\Rightarrow$  the acf is non-summable, i.e., traffic exhibits long range dependence (LRD):  $\sum_k r(k) \rightarrow \infty$

Short-range dependence:  $r(k) \sim \rho^k$ ,  $0 < \rho < 1 \Rightarrow \sum_k r(k) < \infty$ 

X is (exactly) second-order self-similar if:  $r^{(m)}(k) = r(k), k \ge 0$  and (asymptotically) second-order self-similar if:  $r^{(m)}(k) \to r(k), m \to \infty$ 

White noise (not self-similar): 
$$r^{(m)}(k) \rightarrow 0$$
,  $m \rightarrow \infty$ 

### **Auto-correlation**



# Self-Similarity

A traffic trace is self-similar if (equivalently):

4. the expected rescaled adjusted range statistic:

$$\mathbb{E}[R(n)/S(n)] \sim a_5 n^H, n \to \infty, a_5 > 0$$

has Hurst parameter  $\frac{1}{2} < H < 1$ 

The Hurst parameter expresses the speed of decay of the acf

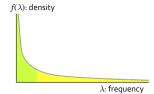
- $H \le \frac{1}{2}$ : short-range dependent processes, e.g., Poisson, batch-Poisson, Markov-modulated Poisson
- H > 1: non-stationary process

# Self-Similarity

A traffic trace is self-similar if (equivalently):

3. taking the time series into the frequency domain (Fourier transform), the low frequency components obeys a power-law near the origin (a low frequency is proportionally denser than its next higher frequency):

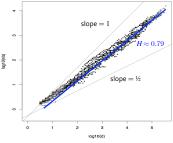
$$\begin{split} f(\lambda) &\sim a_3 \lambda^{-\gamma}, \ \lambda \rightarrow 0, 0 < \gamma < 1, a_3 > 0 \\ \gamma &= 1 - \beta \end{split}$$



Cf. Zipf distribution

## Detecting LRD in Ethernet Trace

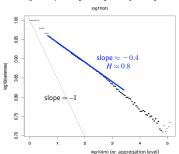
 ${\cal H}$  can estimated directly from R/S statistic



Or from variance time plot

• slope 
$$-\beta$$
,  $0 < \beta < 1$ 

• 
$$H = 1 - \beta/2$$

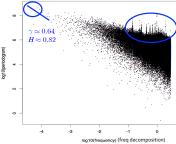


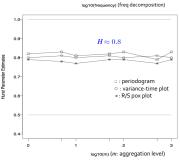
Detecting LRD in Ethernet Trace

Or from periodogram, slope of 10% of the lowest frequencies, near 0

$$H = (1+\gamma)/2$$

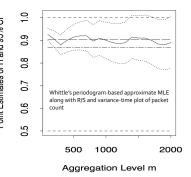
Hurst parameter stays constant across traffic aggregation levels





## Detecting LRD in Ethernet Trace

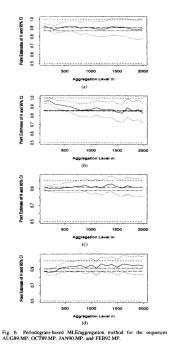
H can also be estimated with maximum-likelihood estimator (MLE) based on the periodogram (Whittle estimator) with the advantage of computing 95% confidence interval



# Detecting LRD in Ethernet Trace

*H* increases as traffic load increases!

Remain true over time (89-92)

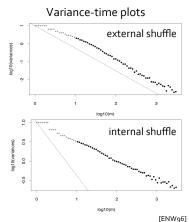


## **Implication**

Long-range dependent traffic effects queueing delay: it makes buffer sizing ineffectual

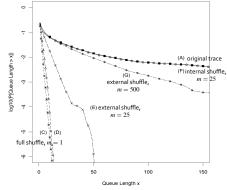
How do we know LRD causes ineffectual buffering [ENW96]?

- external shuffle experiment: divide traffic into *m* blocks and shuffle the blocks around preserving the sequence inside each block: destroys LRD, preserves SRD
- internal shuffle experiment: same blocks, shuffle traffic inside each block, keeping the block order: destroys SRD, preserves LRD



## **Implication**

Resulting queue occupancy statistics:



1-p(ON)

1-p(OFF)

OFF

ON

What can we do about it?

- frequency domain view: traffic can be decomposed into high (spikes), mid (ripples), and low (swells) frequencies
- network must have enough capacity to handle peak rate of low frequency
- buffer space should be used only to handle high-frequency traffic

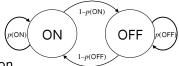
[ENW96]

#### Causes of LRD

Aggregation of ON/OFF traffic with heavy-tailed OFF time distribution [W+97]

- human "think" time
- effect of TCP congestion avoidance (cwnd)
- multimedia sources can also be modeled as ON/OFF

Why does long-tailed ON/OFF distributions cause LRD?



- long OFF time means autocorrelation of bursts at large k, hence  $\sum r(k) \rightarrow \infty$
- long ON time increases the probability of seeing other traffic

# Heavy-tailed Distributions

$$P[X > x] \sim x^{-\alpha}, x \to \infty, 0 < \alpha < 2$$

Examples: Pareto, Weibul, Zipf

Pareto distribution:

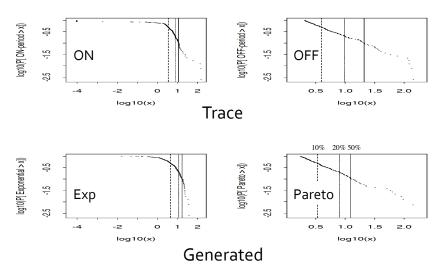
$$p(x) = \alpha k^{\alpha} x^{-\alpha - 1}, \ \alpha \text{ and } k > 0, \ x \ge k$$

$$F(X) = P[X \le x] = 1 - (k/x)^{\alpha}$$

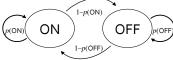
 $\alpha < 2$ : trace has infinite variance

 $\alpha < 1$ : trace has infinite mean

# Heavy-tailed Distributions



## Modeling LRD



Aggregation of ON/OFF traffic with heavy-tailed OFF time distribution [W+97]

Advantage: parsimonious, only one parameter,  $\boldsymbol{\alpha}$ 

$$P[X > x] \sim x^{-\alpha}, x \to \infty, 0 < \alpha < 2$$

#### Alternative models:

- by fitting multiple short-range dependent processes: parameter explosion, no physically meaningful interpretations
- fractional Gaussian noise: does not model short-range dependencies
- fractional ARIMA:
- can model both short- and long-range dependencies,
- but still does not provide physical explanation of self-similarity
- plus, known parameter estimation techniques too expensive

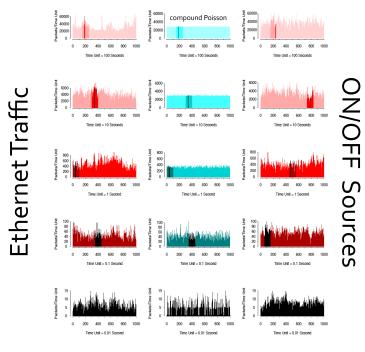
### **Discussions**

Gold standard of measurement study and analysis

Prior to this paper, traffic modeling assumes Poisson distribution

After this paper, traffic modeling uses power law distributions (Pareto, Weibul, Zipf)

A flurry of follow-on papers found power-law distribution everywhere in the network . . .



[W+97]

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Traffic