Probabilistic PCA for Heteroscedastic Data



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Introduction

Weighted PCA

Homoscedastic PPCA (review)

Heterocedastic PPCA: known variances (2019)

Heterocedastic PPCA: unknown variances (2020)

Extensions (2021)



Modern datasets increasingly contain data of varying quality.

For example: mixture of a few high quality (costly) sensors with many low quality (inexpensive) sensors.



Question: How should we account for heterogeneous quality?



Many settings: medical imaging, astronomy, sensor networks, ...



varying radiation levels



varying atmosphere



varying sensor quality

http://www.medicalnewstoday.com/articles/153201.php https://www.nasa.gov/multimedia/imagegallery/iotd.html http://www.livescience.com/27992-portable-pollution-sensors-improve-data-nsf-ria.html



Many settings: medical imaging, astronomy, sensor networks, ...



varying radiation levels (millions of voxels)



varying atmosphere (thousands of pixels)



varying sensor quality (thousands of locations)

Much modern data is high-dimensional...with low-dimensional structure. A standard (well-understood?) tool: Principal Component Analysis

How should we account for heterogeneous quality in PCA?

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```

Preview





HePPCAT: Heteroscedastic probabilistic PCA technique



Model samples $y_1, \ldots, y_n \in \mathbb{R}^d$ as

$$y_i = \mathbf{F} z_i + \sigma_i \varepsilon_i.$$







$$y_i = \mathbf{F} z_i + \sigma_i \varepsilon_i.$$

k latent factors coefficients $d \times k$ matrix

















Samples have <u>heteroscedastic</u> noise. Homoscedastic if $\sigma_1 = \cdots = \sigma_n$.



Recall:

$$y_i = \mathbf{F} z_i + \varepsilon_i \sim \mathcal{N}(\mathbf{F} z_i, \sigma_i^2 \mathbf{I})$$

Negative log-likelihood:

$$L(\boldsymbol{F}, \boldsymbol{Z}) = \sum_{i} \frac{1}{2\sigma_{i}^{2}} \|\boldsymbol{y}_{i} - \boldsymbol{F}\boldsymbol{z}_{i}\|_{2}^{2}.$$

Change of variables: $\tilde{y}_i = y_i / \sigma_i, \ \tilde{z}_i = z_i / \sigma_i$

$$\min_{F, \boldsymbol{X}} \| \tilde{\boldsymbol{Y}} - \boldsymbol{F} \tilde{\boldsymbol{Z}} \|_{\mathrm{F}}^2 \implies \hat{\boldsymbol{F}} = \text{ first } k \text{ left singular vectors of } \tilde{\boldsymbol{Y}}$$

This is "sample-weighted" low-rank approximation / factorization. aka ML factor analysis (Young, 1941) See Hong et al. arXiv 1810.12862 for optimally weighted PCA



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Model: $y_i = \mathbf{F} z_i + \sigma \varepsilon_i \in \mathbb{R}^d$ with $z_i \sim \mathcal{N}(0, \mathbf{I}_k)$ and $\varepsilon_i \sim \mathcal{N}(0, \mathbf{I}_d)$

Review: Probabilistic PCA for homoscedastic noise



Model: $y_i = \mathbf{F} z_i + \sigma \varepsilon_i \in \mathbb{R}^d$ with $z_i \sim \mathcal{N}(0, \mathbf{I}_k)$ and $\varepsilon_i \sim \mathcal{N}(0, \mathbf{I}_d)$ Equivalently, $y_i \sim \mathcal{N}(0, \mathbf{FF}' + \sigma^2 \mathbf{I}_d)$.



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Approach: Maximize the log-likelihood of *F* (dropping constants)

$$\mathcal{L}(\mathbf{F}) \coloneqq \frac{1}{2} \sum_{i=1}^{n} \left\{ \ln \det(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1} - y_{i}'(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1}y_{i} \right\}$$
$$= \frac{n}{2} \left\{ \ln \det(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1} - \frac{1}{n} \sum_{i=1}^{n} y_{i}'(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1}y_{i} \right\}$$
$$= \frac{n}{2} \left[\ln \det(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1} - \operatorname{tr} \left\{ \frac{1}{n} \sum_{i=1}^{n} y_{i}y_{i}'(\mathbf{F}\mathbf{F}' + \sigma^{2}\mathbf{I}_{d})^{-1} \right\} \right].$$

Data appears only through the sample correlation matrix!



$$\mathcal{L}(\boldsymbol{F}) = \frac{n}{2} \left[\ln \det(\boldsymbol{F}\boldsymbol{F}' + \sigma^2 \boldsymbol{I}_d)^{-1} - \operatorname{tr}\left\{ \frac{1}{n} \sum_{i=1}^n y_i y_i' (\boldsymbol{F}\boldsymbol{F}' + \sigma^2 \boldsymbol{I}_d)^{-1} \right\} \right].$$



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Fact (Tipping & Bishop 1999): The likelihood is maximized by

$$\hat{\pmb{F}} = \pmb{V} \operatorname{diag} \Big(\sqrt{\lambda_1 - ar{\lambda}}, \dots, \sqrt{\lambda_k - ar{\lambda}} \Big),$$

where in terms of the sample correlation matrix $\frac{1}{n}\mathbf{Y}\mathbf{Y}'$:

- $V \in \mathbb{R}^{d \times k}$ contains the k principal eigenvectors,
- $\lambda_1, \ldots, \lambda_k$ are the k principal eigenvalues, and
- $\bar{\lambda}$ is the average of the remaining eigenvalues.



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Clean solution: Eigendecomposition + shrinkage



$$\mathcal{L}(\boldsymbol{F}) = \frac{n}{2} \left[\ln \det(\boldsymbol{F}\boldsymbol{F}' + \sigma^2 \boldsymbol{I}_d)^{-1} - \operatorname{tr}\left\{ \frac{1}{n} \sum_{i=1}^n y_i y_i' (\boldsymbol{F}\boldsymbol{F}' + \sigma^2 \boldsymbol{I}_d)^{-1} \right\} \right].$$

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Clean solution: Eigendecomposition + shrinkage Components **V** are same as in ordinary PCA

Setup: $n_1 = 200$ samples with $\sigma_i = 1$, $n_2 = 800$ samples with $\sigma_i = \sigma$. $v_i \sim \mathcal{N}(0, F_*F'_* + \sigma_i^2 I_{100}), \qquad F_* = [\tilde{u}_1, \dots, \tilde{u}_3] \operatorname{diag}(4, 2, 1),$





Heteroscedastic PPCA. D. Hong, K. Gilman, L. Balzano, J. Fessler

Setup: $n_1 = 200$ samples with $\sigma_i = 1$, $n_2 = 800$ samples with $\sigma_i = \sigma$.

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🔤 PPCA: Full data 🛛 💻 PPCA: Group 1





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PPCA: Full data

PPCA: Group 1 = PPCA: Group 2





PPCA degrades as data becomes heteroscedastic. PPCA may work better using only part of the data !? Heteroscedastic PPCA. D. Hong, K. Gilman, L. Balzano, J. Fessler

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Deriving a heterocedastic PPCA



This talk: Develop a heteroscedastic PPCA algorithm. Version 1: Assume noise variances $\{v_i = \sigma_i^2\}$ are known.

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Challenge: Log-likelihood no longer separates so nicely:

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Apparently no closed-form for ML estimate of *F*.

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Apparently no closed-form for ML estimate of *F*.

Approach: Instead, derive an Expectation Maximization (EM) algorithm using the complete data likelihood with complete data $\{(y_i, z_i)\}$

$$\mathcal{L}_{c}(\mathbf{F}) \coloneqq -\sum_{i=1}^{n} \left(\frac{\|y_{i} - \mathbf{F}z_{i}\|_{2}^{2}}{2v_{i}} + \frac{\|z_{i}\|_{2}^{2}}{2} \right).$$

(Tipping & Bishop 1999) derived it for homoscedastic PPCA.

EM for heterocedastic PPCA



$$\mathcal{L}_{c}(\mathbf{F}) := -\sum_{i=1}^{n} \left(\frac{\|y_{i} - \mathbf{F}z_{i}\|_{2}^{2}}{2v_{i}} + \frac{\|z_{i}\|_{2}^{2}}{2} \right).$$

E-Step: Expectation with respect to $z_1, \ldots, z_n | y_1, \ldots, y_n, F_t$ is

$$\bar{\mathcal{L}}(\boldsymbol{F};\boldsymbol{F}_t) = \sum_{i=1}^n \left[\frac{1}{v_i} y_i' \boldsymbol{F} \bar{z}_{t,i} - \frac{1}{2v_i} \operatorname{tr} \{ \boldsymbol{F}' \boldsymbol{F} (\bar{z}_{t,i} \bar{z}_{t,i}' + v_i M_{t,i}) \} \right],$$

up to constants where

$$\mathsf{M}_{t,i} \coloneqq (\mathbf{F}'_t \mathbf{F}_t + v_i \mathbf{I}_k)^{-1}, \qquad \qquad \bar{z}_{t,i} \coloneqq \mathsf{M}_{t,i} \mathbf{F}'_t y_i.$$

M-step: $\bar{\mathcal{L}}(\boldsymbol{F}; \boldsymbol{F}_t)$ is maximized with respect to \boldsymbol{F} by

$$\boldsymbol{F}_{t+1} = \left(\sum_{i=1}^{n} \frac{1}{v_i} y_i \bar{z}'_{t,i}\right) \left(\sum_{i=1}^{n} \frac{1}{v_i} \bar{z}_{t,i} \bar{z}'_{t,i} + M_{t,i}\right)^{-1}.$$

EM for heterocedastic PPCA



$$\mathcal{L}_{c}(\mathbf{F}) := -\sum_{i=1}^{n} \left(\frac{\|y_{i} - \mathbf{F}z_{i}\|_{2}^{2}}{2v_{i}} + \frac{\|z_{i}\|_{2}^{2}}{2} \right).$$

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Now the factor estimates \hat{F} differ from PCA eigenvectors.

Heteroscedastic PPCA effectively uses all data



Setup: $n_1 = 200$ samples with $\sigma_i = 1$, $n_2 = 800$ samples with $\sigma_i = \sigma$.

$$y_i \sim \mathcal{N}(0, \boldsymbol{F}_* \boldsymbol{F}'_* + v_i \boldsymbol{I}_{100}), \qquad \quad \boldsymbol{F}_* = [\tilde{u}_1, \dots, \tilde{u}_3] \operatorname{diag}(4, 2, 1),$$

PPCA: Full data

PPCA: Group 1 PPCA: Group 2

0

0

1

 σ

2

3



3



1

 σ

2

0

0

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🔤 PPCA: Full data 🛛 💻 PPC

💻 PPCA: Group 1 🛛 📒 PPCA: Group 2 📃 HeteroPPCA





But how to initialize for this nonconcave problem?



Setup: $n_1 = 200$ samples with $v_i = 1$, $n_2 = 800$ samples with $v_i = 4$.

$$y_i \sim \mathcal{N}(0, F_*F'_* + v_i I_{100}), \qquad F_* = [\tilde{u}_1, \dots, \tilde{u}_3] \operatorname{diag}(4, 2, 1)^{1/2},$$

Rand init (true eigvals, rand eigvecs) vs. Homoscedastic PPCA init



Interesting phenomenon: iterates do not seem to hit bad local maxima! Why?

What about doing a weighted PCA?



Idea: Give noisier samples less weight in PCA.

Simple tweak: Replace the sample covariance with a *weighted* version

$$\frac{1}{n}\sum_{i=1}^{n}y_{i}y_{i}' \qquad \longrightarrow \qquad \frac{1}{n}\sum_{i=1}^{n}\omega_{i}^{2}y_{i}y_{i}',$$

Choices for weights:

• unweighted: $\omega_i^2 = 1$

• inverse noise variance $\omega_i^2 = 1/\sigma_i^2$

- rescales data to make noise homoscedastic
- corresponds to MLE for low-rank signal $\boldsymbol{F}[z_1, \ldots, z_n]$
- square inverse noise variance $\omega_i^2 = 1/\sigma_i^4$
 - more aggressive down-weighting, can help for low SNRs

High-dim asymptotics studied in (Hong, Fessler, Balzano 2019).

What about doing a weighted PCA? (Jolliffe 2002)





Heteroscedastic PPCA. D. Hong, K. Gilman, L. Balzano, J. Fessler



- PPCA for data with heteroscedastic noise that uses all the data effectively, no matter how noisy
- practical EM algorithm
- interesting phenomenon: do not seem to get stuck in bad local maxima!



- PPCA for data with heteroscedastic noise that uses all the data effectively, no matter how noisy
- practical EM algorithm
- interesting phenomenon: do not seem to get stuck in bad local maxima!
- Ongoing/future: (from Dec. 2019 CAMSAP talk)
 - joint estimation of noise variance
 - optimization landscape
 - other optimization approaches (manifold optimization?, minorize maximize?)
 - analysis of asymptotic performance



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HePPCAT Version 2: unknown noise variances



Model $n_1 + \cdots + n_L = n$ data samples in \mathbb{R}^d from L noise level groups:

$$\mathbf{y}_{\ell,i} = \mathbf{F}\mathbf{z}_{\ell,i} + \boldsymbol{\varepsilon}_{\ell,i}, \quad i \in \{1, \dots, n_\ell\}, \ \ell \in \{1, \dots, L\},$$
(2)

$$\begin{split} \boldsymbol{F} &\in \mathbb{R}^{d \times k} : \text{ deterministic factor matrix to estimate} \\ \boldsymbol{z}_{\ell,i} &\sim \mathcal{N}(\boldsymbol{0}_k, \boldsymbol{I}_k) : \text{ random coefficients} \\ \boldsymbol{\varepsilon}_{\ell,i} &\sim \mathcal{N}(\boldsymbol{0}_d, \boldsymbol{v}_\ell \boldsymbol{I}_d) : \text{ noise vectors} \\ \boldsymbol{v}_1, \dots, \boldsymbol{v}_L : \text{ deterministic noise variances to estimate.} \end{split}$$

Equivalently:

$$oldsymbol{y}_{\ell,i} \sim \mathcal{N}(oldsymbol{0}_d, oldsymbol{F}oldsymbol{F}' + oldsymbol{v}_\elloldsymbol{I}_d)$$



Joint log-likelihood for factors \boldsymbol{F} and variances \boldsymbol{v} (dropping a constant):

$$\mathcal{L}(\boldsymbol{F},\boldsymbol{v}) \triangleq \frac{1}{2} \sum_{\ell=1}^{L} \Big[n_{\ell} \ln \det(\boldsymbol{F}\boldsymbol{F}' + v_{\ell}\boldsymbol{I}_{d})^{-1} - \operatorname{tr} \big\{ \boldsymbol{Y}_{\ell}'(\boldsymbol{F}\boldsymbol{F}' + v_{\ell}\boldsymbol{I}_{d})^{-1} \boldsymbol{Y}_{\ell} \big\} \Big],$$

$$\mathbf{Y}_{\ell} \triangleq [\mathbf{y}_{\ell,1}, \dots, \mathbf{y}_{\ell,n_{\ell}}] \in \mathbb{R}^{d \times n_{\ell}}$$
 for $\ell \in \{1, \dots, L\}$:
data matrices for each of the *L* groups.

Similar to earlier log-likelihood (1).

Apparently no closed form ML estimates for F and vNo closed-form ML estimate of F given vNo closed-form ML estimate of v given F \implies Alternating ascent algorithms

Alternating ascent

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F update : EM update similar to Version 1 HePPCAT **v** update : separates into L univariate maximizations (over $v_{\ell} \ge 0$)

$$\mathcal{L}_{\ell}(\mathbf{v}_{\ell}) \triangleq -\sum_{j=0}^{k} \left\{ \alpha_{j} \ln(\gamma_{j} + \mathbf{v}_{\ell}) + \frac{\beta_{j}}{\gamma_{j} + \mathbf{v}_{\ell}} \right\},$$
(3)

where $\alpha_0 \triangleq d - k$, $\beta_0 \triangleq \| (I_d - U_t U'_t) Y_\ell \|_{\mathrm{F}}^2 / n_\ell$, $\gamma_0 \triangleq 0$,

$$j \ge 1$$
: $\alpha_j \triangleq 1$, $\beta_j \triangleq \| \mathbf{Y}'_{\ell} \boldsymbol{\mu}_{t,j} \|_2^2 / n_{\ell}$, $\gamma_j \triangleq \lambda_{t,j}$,

and $U_t = [\mu_{t,1}, \dots, \mu_{t,k}]$ and $\lambda_t = (\lambda_{t,1}, \dots, \lambda_{t,k})$ are the eigenvectors and eigenvalues of $F_t F'_t$ at iteration t.

Univariate but non-convex. Methods developed and investigated: root finding, EM, minorize-maximize (MM), difference of concave...

Algorithm examples





Figure: Comparison of update methods for $n = 10^3$ samples in $d = 10^2$ dimensions with k = 3 underlying factors $\lambda_* = (4, 2, 1)$. The noise is heteroscedastic: the first $n_1 = 200$ samples have noise variance $\tilde{v}_1 = 1$ and the remaining $n_2 = 800$ have $\tilde{v}_2 = 4$. Walltimes are medians taken over 100 runs of the algorithm to reduce the effect of experimental noise. Markers are placed every five iterations.

Statistical performance example





Figure: Comparison with homoscedastic PPCA applied on: i) full data, ii) only group 1, i.e, the $n_1 = 200$ samples with noise variance $\tilde{v}_1 = 1$, and iii) only group 2, i.e., the $n_2 = 800$ samples with noise variance $\tilde{v}_2 = \sigma_2^2$. Lower is better in (a), and higher is better in (b)-(d). HePPCAT outperforms the homoscedastic methods on all four metrics.

Here, problem size is large enough that HePPCAT with unknown variance.

Effect of block size on variance estimates





Figure: Estimated noise variances for various block sizes.





Figure: Normalized factor estimation error (median and interquartile intervals) for varying block sizes.

The three block sizes lead to practically identical performance here.



10

5

- 0

-5

-10

Single data realization (heatmap). 20Features 1002001000 Samples Normalized error in $\widetilde{\mathrm{U}}\widetilde{\mathrm{U}}' \|_F / \| \widetilde{\mathrm{U}}\widetilde{\mathrm{U}}' \|_F$ 1.2subspace estimation (lower is better). 1.00.8Zhang, Cai, Wu. 2018 0.6Heteroscedastic PPCA ArXiv 1810.08316 ŰŬ(0.42

Heteroscedastic PPCA. D. Hong, K. Gilman, L. Balzano, J. Fessler

 σ_2

3

Apparently favorable convergence





Figure: Convergence gaps of each algorithm to the maximum converged log-likelihood per heteroscedastic noise experiment. n = [200, 800] and $v_1 = 1$. Three different types of initializations.



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Incremental / online algorithms



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- Acceleration via momentum



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- missing data, heterogeneity across features



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- bias correction for variance estimates
- ▶ probabalistic dictionary learning (**F** wide, **z** sparse, so non-normal)

Supervised HePPCAT (new in 2021)



Data $\boldsymbol{Y} \in \mathbb{R}^{d \times n}$; Factors $\boldsymbol{F} \in \mathbb{R}^{d \times k}$; Clay's version:

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^k} \min_{\boldsymbol{F} \in \mathbb{R}^{d \times k}} L(\boldsymbol{Y}' \boldsymbol{F} \boldsymbol{\beta}) + \lambda \| \boldsymbol{Y} - \boldsymbol{F} \boldsymbol{F}' \boldsymbol{Y} \|_{\mathrm{F}}^2.$$
(4)

(Any statistical model associated with norm? If so, then homoscedastic.)



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(Any statistical model associated with norm? If so, then homoscedastic.) Supervised HePPCAT for known noise variances:

$$\min_{\boldsymbol{\beta} \in \mathbb{R}^{k}} \min_{\boldsymbol{F} \in \mathbb{R}^{d \times k}} L(\boldsymbol{Y}'\boldsymbol{F}\boldsymbol{\beta}) + \underbrace{\frac{\lambda}{2} \sum_{i} -\ln \det(\boldsymbol{F}\boldsymbol{F}' + \sigma_{i}^{2}) + \boldsymbol{y}_{i}'(\boldsymbol{F}\boldsymbol{F}' + \sigma_{i}^{2})^{-1}\boldsymbol{y}_{i}}_{\text{neg. log likelihood}}.$$

EM majorizer for F update is quadratic, not "quartic" like (4):

$$L(\mathbf{Y}'\boldsymbol{F}\boldsymbol{\beta}) + \sum_{i} \frac{\lambda}{2\sigma_{i}^{2}} \left[-2\mathbf{y}_{i}'\boldsymbol{F}\bar{\mathbf{z}}_{i} - \bar{\mathbf{z}}_{i}'\boldsymbol{F}'\boldsymbol{F}\bar{\mathbf{z}}_{i} + \operatorname{tr}(\boldsymbol{F}\bar{\boldsymbol{M}}\boldsymbol{F}') \right].$$