

REGULARIZATION

Nonlinear Feature Maps

One way to create nonlinear estimators or classifiers is to first transform the data via a nonlinear feature map

$$\Phi : \mathbb{R}^d \rightarrow \mathcal{H}$$

and apply a linear method to the transformed data $\Phi(x_1), \dots, \Phi(x_n)$

Example 1 $y_i = f(x_i) + \epsilon_i, i=1, \dots, n$

$$f(x) = \sum_{j=0}^p \beta_j x^j \quad (\text{degree } p \text{ polynomial})$$

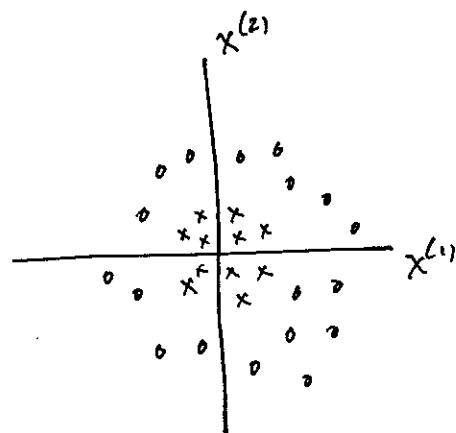
To estimate f , we may apply least-squares regression to the transformed data $\Phi(x_i)$,

where $\Phi(x) = [1 \ x \ x^2 \dots \ x^p]^T$

$$\Rightarrow \hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{y}, \quad \mathbf{A} =$$

(A)

Example 2



$$x \mapsto \Phi(x) = \begin{bmatrix} 1 \\ x^{(1)} \\ x^{(2)} \\ x^{(1)}, x^{(2)} \\ (x^{(1)})^2 \\ (x^{(2)})^2 \end{bmatrix}$$

Then the data are linearly separable in the new feature space. They are correctly classified by

$$\text{sign}\{w^T \Phi(x)\}$$

where

(B)

$$w =$$

In many applications, we don't know exactly how to design $\Phi(x)$, we just know that some nonlinear features are probably important.

In such situations, it is common to include a large number of nonlinear features, in hopes that some of them are relevant.

Unfortunately, this practice can lead to ————— problems.

Example 1, revisited | In least squares polynomial regression, we have

$$\hat{\beta} = (A^T A)^{-1} A^T \underline{y}$$

where

$$A = \begin{bmatrix} 1 & x_1 & x_1^2 & \dots & x_1^p \\ 1 & x_2 & x_2^2 & \dots & x_2^p \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_n & x_n^2 & \dots & x_n^p \end{bmatrix}$$

As p increases, the smallest eigenvalue of $A^T A$ gets very, very small, so that matrix

④ inversion is numerically _____.

Alternatively, the least-squares criterion becomes very _____ near the minimizer.

Essentially, when there are many features, it is possible that a huge coefficient on one feature could be cancelled out by other features.

To remedy the situation, we can incorporate a regularization term into design criteria that will keep coefficients small.

Owing to computational convenience, the most common kind of regularization is _____. We'll examine two cases in detail.

Ridge Regression

Given $y_i = f(x_i) + \epsilon_i$, where $f(x_i) = \beta^T x_i + \beta_0$.

Instead of minimizing the sum of squared errors, in ridge regression, we minimize

$$\sum_{i=1}^n (y_i - \beta^T x_i - \beta_0)^2 + \lambda \|\beta\|^2$$

where $\lambda > 0$ is a tuning parameter.

Note: β_0 is not penalized so that our solution is independent of where the origin is placed.

Let's derive the solution. First, let's eliminate β_0 .

$$\frac{\partial}{\partial \beta_0} = -2 \sum_{i=1}^n (y_i - \beta^T x_i - \beta_0) = 0$$

$$\Rightarrow \hat{\beta}_0 = \frac{1}{n} \sum_i y_i - \hat{\beta}^T x_i$$

(E)

=

Thus, we are left to minimize

$$\sum_{i=1}^n (y_i - \bar{y} - \beta^T(x_i - \bar{x}))^2 + \lambda \beta^T \beta$$

wrt β . For convenience, assume $\bar{y} = 0$, $\bar{x} = 0$.

The criterion may be written

$$(\underline{y} - A\beta)^T(\underline{y} - A\beta) + \lambda \beta^T \beta, \quad A =$$

$$\begin{bmatrix} x_1^{(1)} & \dots & x_1^{(\lambda)} \\ \vdots & & \vdots \\ x_n^{(1)} & \dots & x_n^{(\lambda)} \end{bmatrix}$$

$$= \underline{y}^T \underline{y} + \beta^T A^T A \beta - 2\beta^T A^T \underline{y} + \lambda \beta^T \beta$$

$$= \beta^T [A^T A + \lambda I] \beta - 2\beta^T A^T \underline{y} + \underline{y}^T \underline{y}$$

$$\frac{\partial}{\partial \beta} = 0 \Rightarrow (A^T A + \lambda I) \beta = A^T \underline{y}$$

(F)

\Rightarrow

Observations • $\lambda = 0$ recovers least-squares linear reg.

- λI increases the eigenvalues of $A^T A$ by λ , so that $A^T A + \lambda I$ is not ill-conditioned.

Soft Margin Hyperplane

The training error of a linear classifier may be bounded as

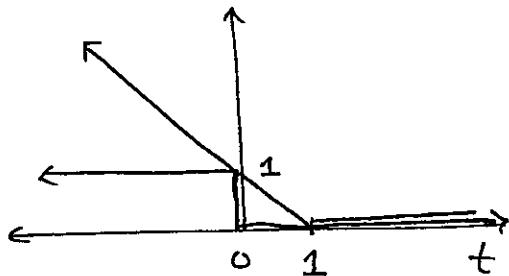
$$\frac{1}{n} \sum_{i=1}^n \mathbb{1}_{\{y_i(\omega^T x_i + b) < 0\}}$$

$$\leq \frac{1}{n} \sum \phi(y_i(\omega^T x_i + b))$$

where $\phi(t)$ is any upper bound on $\mathbb{1}_{\{t < 0\}}$.

Let's take

$$\begin{aligned}\phi(t) &= \max\{0, 1-t\} \\ &=: (1-t)_+\end{aligned}$$



In addition, let's add a quadratic penalty to keep the coefficients small.

$$\Rightarrow \min_{w, b} \frac{1}{2} \|w\|^2 + \frac{1}{n} \sum_{i=1}^n (1 - y_i(\omega^T x_i + b))_+$$

Compare this to the quadratic program for the optimal soft-margin hyperplane:

$$\begin{aligned} \min_{w, b, \xi_i} \quad & \frac{1}{2} \|w\|^2 + \frac{C}{n} \sum_{i=1}^n \xi_i \\ \text{s.t.} \quad & y_i (w^T x_i + b) \geq 1 - \xi_i, \quad i=1, \dots, n \\ & \xi_i \geq 0, \quad i=1, \dots, n \end{aligned}$$

Claim] If $C = \frac{1}{\lambda}$, these two optimization problems are solved by the same w, b .

(G)

Proof:

Key

A.

$$A = \begin{bmatrix} \vdots & \overline{\mathbb{I}}(x_1)^T \\ \vdots & \overline{\mathbb{I}}(x_2)^T \\ \vdots & \vdots \\ \vdots & \overline{\mathbb{I}}(x_n)^T \end{bmatrix} = \begin{bmatrix} 1 & x_1 & x_1^2 & \cdots & x_1^P \\ 1 & x_2 & x_2^2 & \cdots & x_2^P \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_n & x_n^2 & \cdots & x_n^P \end{bmatrix}$$

B.

$$w = \begin{bmatrix} -r^2 \\ 0 \\ 0 \\ 0 \\ \vdots \\ 1 \end{bmatrix}$$

(circle of radius r)

C. ill-conditioned

D. unstable, flat, quadratic

E. $\bar{y} - \hat{\beta}^T \bar{x}$ F. $\hat{\beta} = (A^T A + \lambda I)^{-1} A^T \bar{y}$

G. Since scaling an objective function by a positive constant does not change the solution, it suffices to show that

$$\text{QP1} \quad \min_{(w,b)} \frac{1}{2} \|w\|^2 + \frac{1}{n\lambda} \sum (1 - y_i(w^T x_i + b))_+$$

QP2 $\min_{(w,b,\xi)} \frac{1}{2} \|w\|^2 + \frac{1}{n\lambda} \sum \xi_i$

s.t. $y_i(w^T x_i + b) \geq 1 - \xi_i, \xi_i \geq 0$

are solved by the same (w, b) .

- Suppose (w^*, b^*) is an optimizer of QP1

Claim : (w^*, b^*, ξ^*) is an optimizer of QP2, where

$$\xi_i^* = \max(0, 1 - y_i(w^T x_i + b^*))$$

Suppose not, and let (w, b, ξ) be an optimizer of QP2.

Since (w, b, ξ) is a global optimizer

- If $\xi_i > 0$, then $y_i(w^T x_i + b) = 1 - \xi_i$

[otherwise, we could decrease the objective function without violating the constraints]

- If $\xi_i = 0$, then $y_i(w^T x_i + b) \geq 1$

Thus $\sum \xi_i = \sum (1 - y_i(w^T x_i + b))_+$ and so

$$\begin{aligned} & \frac{1}{2} \|w\|^2 + \frac{1}{n\lambda} \sum (1 - y_i(w^T x_i + b))_+ \\ &= \frac{1}{2} \|w\|^2 + \frac{1}{n\lambda} \sum \xi_i \\ &< \frac{1}{2} \|w^*\|^2 + \frac{1}{n\lambda} \sum \xi_i^* \\ &= \frac{1}{2} \|w^*\|^2 + \frac{1}{n\lambda} \sum (1 - y_i(w^{*T} x_i + b))_+ \end{aligned}$$

which contradicts optimality of (w^*, b^*) for QP1.

- Suppose (w^*, b^*, ξ^*) is an optimizer of QP2.

Claim : (w^*, b^*) is an optimizer of QP1.

The argument is similar and is left as an exercise.